

CHAPTER 5

Section 5-1

5-1. First, $f(x,y) \geq 0$. Let R denote the range of (X,Y).

Then, $\sum_R f(x,y) = \frac{1}{4} + \frac{1}{8} + \frac{1}{4} + \frac{1}{4} + \frac{1}{8} = 1$

a) $P(X < 2.5, Y < 3) = f(1.5, 2) + f(1, 1) = 1/8 + 1/4 = 3/8$

b) $P(X < 2.5) = f(1.5, 2) + f(1.5, 3) + f(1, 1) = 1/8 + 1/4 + 1/4 = 5/8$

c) $P(Y < 3) = f(1.5, 2) + f(1, 1) = 1/8 + 1/4 = 3/8$

d) $P(X > 1.8, Y > 4.7) = f(3, 5) = 1/8$

e) $E(X) = 1(1/4) + 1.5(3/8) + 2.5(1/4) + 3(1/8) = 1.8125$

$E(Y) = 1(1/4) + 2(1/8) + 3(1/4) + 4(1/4) + 5(1/8) = 2.875$

$V(X) = E(X^2) - [E(X)]^2 = [1^2(1/4) + 1.5^2(3/8) + 2.5^2(1/4) + 3^2(1/8)] - 1.8125^2 = 0.4961$

$V(Y) = E(Y^2) - [E(Y)]^2 = [1^2(1/4) + 2^2(1/8) + 3^2(1/4) + 4^2(1/4) + 5^2(1/8)] - 2.875^2 = 1.8594$

f) marginal distribution of X

x	f(x)
1	1/4
1.5	3/8
2.5	1/4
3	1/8

g) $f_{Y|2.5}(y) = \frac{f_{XY}(2.5, y)}{f_X(2.5)}$ and $f_X(2.5) = 1/4$. Then,

y	$f_{Y 2.5}(y)$
4	(1/4)/(1/4)=1

h) $f_{X|2}(x) = \frac{f_{XY}(x, 2)}{f_Y(2)}$ and $f_Y(2) = 1/8$. Then,

x	$f_{X 2}(y)$
1.5	(1/8)/(1/8)=1

i) $E(Y|X=1.5) = 2(1/3) + 3(2/3) = 2/1/3$

j) Since $f_{Y|1.5}(y) \neq f_Y(y)$, X and Y are not independent

5-2. Let R denote the range of (X,Y). Because

$\sum_R f(x, y) = c(2+3+4+3+4+5+4+5+6) = 1$, $36c = 1$, and $c = 1/36$

a) $P(X=1, Y < 3) = f_{XY}(1, 1) + f_{XY}(1, 2) = \frac{1}{36}(2+3) = 5/36$

b) $P(X=1)$ is the same as part (a) = 1/4

c) $P(Y=2) = f_{XY}(1, 2) + f_{XY}(2, 2) + f_{XY}(3, 2) = \frac{1}{36}(3+4+5) = 1/3$

d) $P(X < 2, Y < 2) = f_{XY}(1, 1) = \frac{1}{36}(2) = 1/18$

e)

$$E(X) = 1[f_{XY}(1, 1) + f_{XY}(1, 2) + f_{XY}(1, 3)] + 2[f_{XY}(2, 1) + f_{XY}(2, 2) + f_{XY}(2, 3)]$$

$$+ 3[f_{XY}(3, 1) + f_{XY}(3, 2) + f_{XY}(3, 3)]$$

$$= (1 \times \frac{9}{36}) + (2 \times \frac{12}{36}) + (3 \times \frac{15}{36}) = 13/6 = 2.167$$

$$V(X) = (1 - \frac{13}{6})^2 \cdot \frac{9}{36} + (2 - \frac{13}{6})^2 \cdot \frac{12}{36} + (3 - \frac{13}{6})^2 \cdot \frac{15}{36} = 0.639$$

$$E(Y) = 2.167$$

$$V(Y) = 0.639$$

f) Marginal distribution of X

x	$f_X(x) = f_{XY}(x,1) + f_{XY}(x,2) + f_{XY}(x,3)$
1	1/4
2	1/3
3	5/12

$$\text{g) } f_{Y|X}(y) = \frac{f_{XY}(1,y)}{f_X(1)}$$

y	$f_{Y X}(y)$
1	$(2/36)/(1/4)=2/9$
2	$(3/36)/(1/4)=1/3$
3	$(4/36)/(1/4)=4/9$

$$\text{h) } f_{X|Y}(x) = \frac{f_{XY}(x,2)}{f_Y(2)} \text{ and } f_Y(2) = f_{XY}(1,2) + f_{XY}(2,2) + f_{XY}(3,2) = \frac{12}{36} = 1/3$$

x	$f_{X Y}(x)$
1	$(3/36)/(1/3)=1/4$
2	$(4/36)/(1/3)=1/3$
3	$(5/36)/(1/3)=5/12$

$$\text{i) } E(Y|X=1) = 1(2/9) + 2(1/3) + 3(4/9) = 20/9$$

j) Since $f_{XY}(x,y) \neq f_X(x)f_Y(y)$, X and Y are not independent.

5-3. $f(x, y) \geq 0$ and $\sum_K f(x, y) = 1$

$$\text{a) } P(X < 0.5, Y < 1.5) = f_{XY}(-1, -2) + f_{XY}(-0.5, -1) = \frac{1}{8} + \frac{1}{4} = \frac{3}{8}$$

$$\text{b) } P(X < 0.5) = f_{XY}(-1, -2) + f_{XY}(-0.5, -1) = \frac{3}{8}$$

$$\text{c) } P(Y < 1.5) = f_{XY}(-1, -2) + f_{XY}(-0.5, -1) + f_{XY}(0.5, 1) = \frac{7}{8}$$

$$\text{d) } P(X > 0.25, Y < 4.5) = f_{XY}(0.5, 1) + f_{XY}(1, 2) = \frac{5}{8}$$

$$\text{e) } E(X) = -1(\frac{1}{8}) - 0.5(\frac{1}{4}) + 0.5(\frac{1}{2}) + 1(\frac{1}{8}) = \frac{1}{8}$$

$$E(Y) = -2(\frac{1}{8}) - 1(\frac{1}{4}) + 1(\frac{1}{2}) + 2(\frac{1}{8}) = \frac{1}{4}$$

$$V(X) = (-1-1/8)^2(1/8) + (-0.5-1/8)^2(1/4) + (0.5-1/8)^2(1/2) + (1-1/8)^2(1/8) = 0.4219$$

$$V(Y) = (-2-1/4)^2(1/8) + (-1-1/4)^2(1/4) + (1-1/4)^2(1/2) + (2-1/4)^2(1/8) = 1.6875$$

f) marginal distribution of X

x	$f_X(x)$
-1	1/8
-0.5	1/4
0.5	1/2
1	1/8

g) $f_{Y|X}(y) = \frac{f_{XY}(0.5, y)}{f_X(0.5)}$

y	$f_{Y X}(y)$
1	$1/2/(1/2)=1$

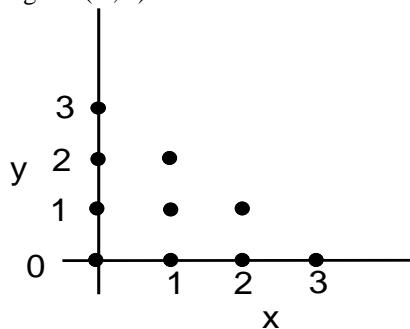
h) $f_{X|Y}(x) = \frac{f_{XY}(x, 1)}{f_Y(1)}$

x	$f_{X Y}(x)$
0.5	$1/2/(1/2)=1$

- i) $E(X|Y=1) = 0.5$
j) No, X and Y are not independent

5-4. Because X and Y denote the number of printers in each category,
 $X \geq 0, Y \geq 0$ and $X + Y = 5$

5-5. a) The range of (X, Y) is



x,y	$f_{xy}(x,y)$
0,0	0.857375
0,1	0.1083
0,2	0.00456
0,3	0.000064
1,0	0.027075
1,1	0.00228
1,2	0.000048
2,0	0.000285
2,1	0.000012
3,0	0.000001

b)

x	$f_x(x)$
0	0.970299
1	0.029403
2	0.000297
3	0.000001

c) $E(X) = 0(0.970299) + 1(0.029403) + 2(0.000297) + 3(0.000001) = 0.03$
or $np = 3(0.01) = 0.03$

d) $f_{Y|2}(y) = \frac{f_{XY}(2,y)}{f_X(2)}$, $f_X(2) = 0.000297$

y	$f_{Y 1}(x)$
0	0.9596
1	0.0404

e) $E(Y|X=1) = 0(0.9596) + 1(0.0404) = 0.0404$

g) No, X and Y are not independent because, for example, $f_Y(0) \neq f_{Y|1}(0)$.

- 5-6. a) The range of (X,Y) is $X \geq 0, Y \geq 0$ and $X+Y \leq 4$. Here X is the number of pages with moderate graphic content and Y is the number of pages with high graphic output among a sample of 4 pages.

The following table is for sampling without replacement. Students would have to extend the hypergeometric distribution to the case of three classes (low, moderate, and high).

For example, $P(X=1, Y=2)$ is calculated as

$$P(X=1, Y=2) = \frac{\binom{60}{1} \binom{30}{1} \binom{10}{2}}{\binom{100}{4}} = \frac{60(30)(45)}{100(99)(98)(97)} = 0.02066$$

	x=0	x=1	x=2	x=3	x=4
y=4	5.35×10^{-5}	0	0	0	0
y=3	0.00184	0.00092	0	0	0
y=2	0.02031	0.02066	0.00499	0	0
y=1	0.08727	0.13542	0.06656	0.01035	0
y=0	0.12436	0.26181	0.19635	0.06212	0.00699

b)

	$f(y)$
y=4	0.0000535
y=3	0.0027600
y=2	0.0459600
y=1	0.2996000
y=0	0.6516300

c) $E(Y) =$

$$\sum_0^4 y_i f(y_i) = 0(0.65163) + 1(0.2996) + 2(0.04596) + 3(0.00276) + 4(0.0000535) = 0.400014$$

d) $f_{Y|3}(y) = \frac{f_{XY}(3,y)}{f_X(3)}$, $f_X(3) = 0.0725$

y	$f_{Y 3}(y)$
0	0.857
1	0.143

2	0
3	0
4	0

- e) $E(Y|X=3) = 0(0.857) + 1(0.143) = 0.143$
 f) $V(Y|X=3) = 0^2(0.857) + 1^2(0.143) - 0.143^2 = 0.123$
 g) No, X and Y are not independent

- 5-7. a) The range of (X,Y) is $X \geq 0, Y \geq 0$ and $X+Y \leq 4$.

Here X and Y denote the number of defective items found with inspection device 1 and 2, respectively.

$$f(x,y) = \left[\binom{4}{x} (0.994)^x (0.006)^{4-x} \right] \left[\binom{4}{y} (0.997)^y (0.003)^{4-y} \right]$$

For $x = 0, 1, 2, 3, 4$ and $y = 0, 1, 2, 3, 4$

	$x = 0$	$x = 1$	$x = 2$	$x = 3$	$x = 4$
$y = 0$	1.05×10^{-19}	6.96×10^{-17}	1.73×10^{-14}	1.91×10^{-12}	7.91×10^{-11}
$y = 1$	1.4×10^{-16}	9.28×10^{-14}	2.305×10^{-11}	2.55×10^{-9}	1.054×10^{-7}
$y = 2$	6.96×10^{-14}	4.61×10^{-11}	1.15×10^{-8}	1.27×10^{-7}	5.24×10^{-5}
$y = 3$	1.54×10^{-11}	1.02×10^{-8}	2.54×10^{-6}	2.81×10^{-4}	0.0116
$y = 4$	1.28×10^{-9}	8.49×10^{-7}	2.11×10^{-4}	0.0233	0.965

$$f(x,y) = \left[\binom{4}{x} (0.993)^x (0.007)^{4-x} \right] \left[\binom{4}{y} (0.997)^y (0.003)^{4-y} \right]$$

$$\text{b) } f(x) = \left[\binom{4}{x} (0.994)^x (0.006)^{4-x} \right] \text{ for } x = 1, 2, 3, 4$$

	$x = 0$	$x = 1$	$x = 2$	$x = 3$	$x = 4$
$f(x)$	1.296×10^{-9}	8.59×10^{-7}	2.134×10^{-4}	0.0236	0.9762

- c) Because X has a binomial distribution $E(X) = n(p) = 4 \times (0.994) = 3.976$

$$\text{d) } f_{Y|2}(y) = \frac{f_{XY}(2,y)}{f_X(2)} = f(y), f_X(2) = 2.134 \times 10^{-4}$$

<u>y</u>	<u>$f_{Y 1}(y)=f(y)$</u>
0	8.1×10^{-11}
1	1.08×10^{-7}
2	5.37×10^{-5}
3	0.0119
4	0.988

- e) $E(Y|X=2) = E(Y) = 4(0.997) = 3.988$
 f) $V(Y|X=2) = V(Y) = n(p)(1-p) = 4(0.997)(0.003) = 0.0120$
 g) Yes, X and Y are independent.

- 5-8. a) $P(X = 2) = f_{XYZ}(2,1,1) + f_{XYZ}(2,1,2) + f_{XYZ}(2,2,1) + f_{XYZ}(2,2,2) = 0.5$
 b) $P(X = 1, Y = 2) = f_{XYZ}(1,2,1) + f_{XYZ}(1,2,2) = 0.35$
 c) $P(Z < 1.5) = f_{XYZ}(1,1,1) + f_{XYZ}(1,2,1) + f_{XYZ}(2,1,1) + f_{XYZ}(2,2,1) = 0.44$

d)

$$P(X = 1 \text{ or } Z = 2) = P(X = 1) + P(Z = 2) - P(X = 1, Z = 2) = 0.5 + 0.56 - 0.3 = 0.76$$

e) $E(X) = 1(0.5) + 2(0.5) = 1.5$

f) $P(X = 1 | Y = 1) = \frac{P(X = 1, Y = 1)}{P(Y = 1)} = \frac{0.05 + 0.10}{0.21 + 0.2 + 0.1 + 0.05} = 0.27$

g) $P(X = 1, Y = 1 | Z = 2) = \frac{P(X = 1, Y = 1, Z = 2)}{P(Z = 2)} = \frac{0.1}{0.1 + 0.2 + 0.21 + 0.05} = 0.19$

h) $P(X = 1 | Y = 1, Z = 2) = \frac{P(X = 1, Y = 1, Z = 2)}{P(Y = 1, Z = 2)} = \frac{0.10}{0.10 + 0.21} = 0.32$

i) $f_{X|YZ}(x) = \frac{f_{XYZ}(x, 1, 2)}{f_{YZ}(1, 2)}$ and $f_{YZ}(1, 2) = f_{XYZ}(1, 1, 2) + f_{XYZ}(2, 1, 2) = 0.31$

x	$f_{X YZ}(x)$
1	0.10/0.31=0.32
2	0.21/0.31=0.68

5-9. (a) $f_{XY}(x, y) = (10\%)^x (30\%)^y (60\%)^{4-x-y}$, for $X+Y \leq 4$

$f_{XY}(x, y)$	x	y
0.1296	0	0
0.0648	0	1
0.0324	0	2
0.0162	0	3
0.0081	0	4
0.0216	1	0
0.0108	1	1
0.0054	1	2
0.0027	1	3
0.0036	2	0
0.0018	2	1
0.0009	2	2
0.0006	3	0
0.0003	3	1
0.0001	4	0

(b) $f_X(x) = P(X=x) = \sum_{X+Y \leq 4} f_{XY}(x, y) .$

$f_X(x)$	x
0.2511	0
0.0405	1
0.0063	2
0.0009	3
0.0001	4

(c) $E(X) = \sum xf_X(x) = 0*0.2511 + 1*0.0405 + 2*0.0063 + 3*0.0009 + 4*0.0001 = 0.0562$

(d) $f(y|X=3) = P(Y=y, X=3)/P(X=3)$

$$P(Y=0, X=3) = C^{30}_1 C^5_3 / C^{50}_4 \quad P(Y=1, X=3) = C^{15}_1 C^5_3 / C^{50}_4$$

$P(X=3) = C^{45}_1 C^5_3 / C^{50}_4$, from the hypergeometric distribution with $N=50$, $n=4$, $k=4$, $x=3$

Therefore

$$f(0|X=3) = [C^{30}_1 C^5_3 / C^{50}_4] / [C^{45}_1 C^5_3 / C^{50}_4] = C^{30}_1 / C^{45}_1 = \frac{30}{45} = \frac{2}{3}$$

$$f(1|X=3) = [C^{15}_1 C^5_3 / C^{50}_4] / [C^{45}_1 C^5_3 / C^{50}_4] = \frac{C^{15}_1}{C^{45}_1} = \frac{15}{45} = \frac{1}{3}$$

$f_{Y 3}(y)$	y	x
2/3	0	3
1/3	1	3
0	2	3
0	3	3
0	4	3

(e) $E(Y|X=3) = 0(0.6667) + 1(0.3333) = 0.3333$

(f) $V(Y|X=3) = (0-0.3333)^2(0.6667) + (1-0.3333)^2(0.3333) = 0.0741$

(g) $f_X(0) = 0.2511$, $f_Y(0) = 0.1555$, $f_X(0) * f_Y(0) = 0.039046 \neq f_{XY}(0,0) = 0.1296$
X and Y are not independent.

- 5-10. (a) $P(X < 5) = 0.44 + 0.04 = 0.48$
 (b) $E(X) = 0.43(23) + 0.44(4.2) + 0.04(11.4) + 0.05(130) + 0.04(0) = 18.694$
 (c) $P_{X|Y=0}(X) = P(X = x, Y = 0)/P(Y = 0) = 0.04/0.08 = 0.5$ for $x = 0$ and 11.4
 (d) $P(X>10|Y=0) = P(X=11.4|Y=0) = 0.5$
 (e) $E(X|Y = 0) = 11.4(0.5) + 0(0.5) = 5.7$

- 5-11. (a) $f_{XYZ}(x,y,z)$

$f_{XYZ}(x,y,z)$	Selects(X)	Updates(Y)	Inserts(Z)
0.43	23	11	12
0.44	4.2	3	1
0.04	11.4	0	0
0.05	130	120	0
0.04	0	0	0

(b) $P_{XZ|Y=0}$

$P_{XZ Y=0}(x,y)$	Selects(X)	updates(Y)	Inserts(Z)
4/8 = 0.5	11.4	0	0
4/8 = 0.5	0	0	0

(c) $P(X<6, Y<6|Z = 0) = P(X = 0, Y = 0) = 0.3077$

(d) $E(X|Y = 0, Z = 0) = 0.5(11.4) + 0.5(0) = 5.7$ where this conditional distribution for X was determined in the previous exercise

- 5-12. Let X, Y, and Z denote the number of bits with high, moderate, and low distortion. Then, the joint distribution of X, Y, and Z is multinomial with n=3 and

$p_1 = 0.02$, $p_2 = 0.03$, and $p_3 = 0.95$.

a)

$$P(X = 2, Y = 1) = P(X = 2, Y = 1, Z = 0)$$

$$= \frac{3!}{2!1!0!} (0.02)^2 (0.03)^1 0.95^0 = 3.6 \times 10^{-5}$$

$$\text{b) } P(X = 0, Y = 0, Z = 3) = \frac{3!}{0!0!3!} (0.02)^0 (0.03)^0 0.95^3 = 0.8574$$

c) X has a binomial distribution with n = 3 and p = 0.02. Then, E(X) = 3(0.02) = 0.06 and V(X) = 3(0.02)(0.98) = 0.0588.

d) First find $P(X | Y = 2)$

$$P(Y = 2) = P(X = 1, Y = 2, Z = 0) + P(X = 0, Y = 2, Z = 1)$$

$$= \frac{3!}{1!2!0!} 0.02(0.03)^2 0.95^0 + \frac{3!}{0!2!1!} (0.02)^0 (0.03)^2 0.95^1 = 0.0026$$

$$P(X = 0 | Y = 2) = \frac{P(X = 0, Y = 2)}{P(Y = 2)} = \left(\frac{3!}{0!2!1!} (0.02)^0 (0.03)^2 0.95^1 \right) / 0.0026 = 0.98654$$

$$P(X = 1 | Y = 2) = \frac{P(X = 1, Y = 2)}{P(Y = 2)} = \left(\frac{3!}{1!2!1!} (0.02)^1 (0.03)^2 0.95^0 \right) / 0.0026 = 0.02077$$

$$E(X | Y = 2) = 0(0.98654) + 1(0.02077) = 0.02077$$

$$V(X | Y = 2) = E(X^2) - (E(X))^2 = 0.02077 - (0.02077)^2 = 0.02034$$

5-13. Determine c such that $c \int_0^3 \int_0^3 xy dx dy = c \int_0^3 y \frac{x^2}{2} \Big|_0^3 dy = c(4.5 \frac{y^2}{2} \Big|_0^3) = \frac{81}{4}c$.

Therefore, $c = 4/81$.

$$\text{a) } P(X < 2, Y < 3) = \frac{4}{81} \int_0^3 \int_0^2 xy dx dy = \frac{4}{81} (2) \int_0^3 y dy = \frac{4}{81} (2) \left(\frac{9}{2}\right) = 0.4444$$

b) $P(X < 2.5) = P(X < 2.5, Y < 3)$ because the range of Y is from 0 to 3.

$$P(X < 2.5, Y < 3) = \frac{4}{81} \int_0^3 \int_0^{2.5} xy dx dy = \frac{4}{81} (3.125) \int_0^3 y dy = \frac{4}{81} (3.125) \frac{9}{2} = 0.6944$$

$$\text{c) } P(1 < Y < 2.5) = \frac{4}{81} \int_1^3 \int_0^{2.5} xy dx dy = \frac{4}{81} (4.5) \int_1^{2.5} y dy = \frac{18}{81} \frac{y^2}{2} \Big|_1^{2.5} = 0.5833$$

$$\text{d) } P(X > 1.9, 1 < Y < 2.5) = \frac{4}{81} \int_{1.9}^{2.5} \int_0^3 xy dx dy = \frac{4}{81} (2.7) \int_1^{2.5} y dy = \frac{4}{81} (2.7) \frac{(2.5^2 - 1)}{2} = 0.35$$

$$\text{e) } E(X) = \frac{4}{81} \int_0^3 \int_0^3 x^2 y dx dy = \frac{4}{81} \int_0^3 9 y dy = \frac{4}{9} \frac{y^2}{2} \Big|_0^3 = 2$$

$$\text{f) } P(X < 0, Y < 4) = \frac{4}{81} \int_0^4 \int_0^0 xy dx dy = 0 \int_0^4 y dy = 0$$

$$\text{g) } f_X(x) = \int_0^3 f_{XY}(x, y) dy = x \frac{4}{81} \int_0^3 y dy = \frac{4}{81} x (4.5) = \frac{2x}{9} \quad \text{for } 0 < x < 3.$$

h) $f_{Y|1.5}(y) = \frac{f_{XY}(1.5, y)}{f_X(1.5)} = \frac{\frac{4}{81}y(1.5)}{\frac{2}{9}(1.5)} = \frac{2}{9}y \quad \text{for } 0 < y < 3.$

i) $E(Y|X=1.5) = \int_0^3 y \left(\frac{2}{9}y \right) dy = \frac{2}{9} \int_0^3 y^2 dy = \frac{2y^3}{27} \Big|_0^3 = 2$

j) $P(Y < 2 | X = 1.5) = f_{Y|1.5}(y) = \int_0^2 \frac{2}{9}y dy = \frac{1}{9}y^2 \Big|_0^2 = \frac{4}{9} - 0 = \frac{4}{9}$

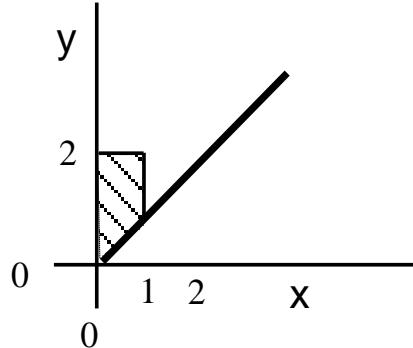
k) $f_{X|2}(x) = \frac{f_{XY}(x, 2)}{f_Y(2)} = \frac{\frac{4}{81}x(2)}{\frac{2}{9}(2)} = \frac{2}{9}x \quad \text{for } 0 < x < 3.$

5-14.

$$\begin{aligned} c \int_0^3 \int_x^{x+2} (x+y) dy dx &= \int_0^3 xy + \frac{y^2}{2} \Big|_x^{x+2} dx \\ &= \int_0^3 \left[x(x+2) + \frac{(x+2)^2}{2} - x^2 - \frac{x^2}{2} \right] dx \\ &= c \int_0^3 (4x+2) dx = [2x^2 + 2x]_0^3 = 24c \end{aligned}$$

Therefore, $c = 1/24$.

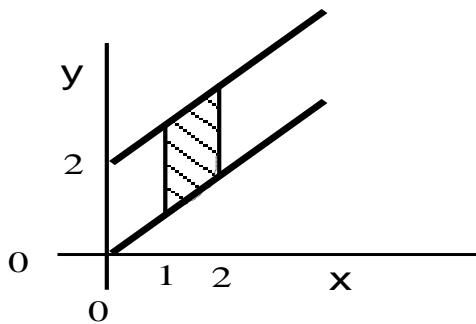
a) $P(X < 1, Y < 2)$ equals the integral of $f_{XY}(x, y)$ over the following region.



Then,

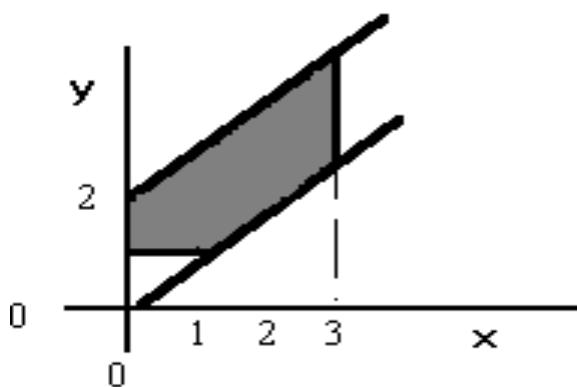
$$\begin{aligned} P(X < 1, Y < 2) &= \frac{1}{24} \int_0^1 \int_x^2 (x+y) dy dx = \frac{1}{24} \int_0^1 xy + \frac{y^2}{2} \Big|_x^2 dx = \frac{1}{24} \int_0^1 2x + 2 - \frac{3x^2}{2} dx = \\ &= \frac{1}{24} \left[x^2 + 2x - \frac{x^3}{2} \Big|_0^1 \right] = 0.10417 \end{aligned}$$

b) $P(1 < X < 2)$ equals the integral of $f_{XY}(x, y)$ over the following region.



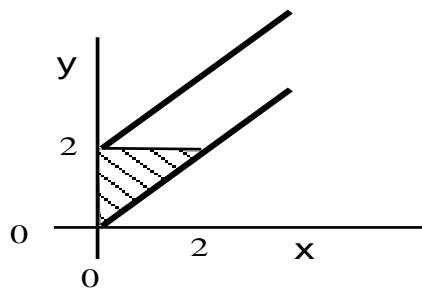
$$\begin{aligned}
 P(1 < X < 2) &= \frac{1}{24} \int_1^2 \int_x^{x+2} (x+y) dy dx = \frac{1}{24} \int_1^2 xy + \frac{y^2}{2} \Big|_x^{x+2} dx \\
 &= \frac{1}{24} \int_1^3 (4x+2) dx = \frac{1}{24} \left[2x^2 + 2x \Big|_1^3 \right] = \frac{1}{6}.
 \end{aligned}$$

c) $P(Y > 1)$ is the integral of $f_{XY}(x, y)$ over the following region.



$$\begin{aligned}
 P(Y > 1) &= 1 - P(Y \leq 1) = 1 - \frac{1}{24} \int_0^1 \int_0^1 (x+y) dy dx = 1 - \frac{1}{24} \int_0^1 \left(xy + \frac{y^2}{2} \right) \Big|_0^1 \\
 &= 1 - \frac{1}{24} \int_0^1 x + \frac{1}{2} - \frac{3}{2} x^2 dx = 1 - \frac{1}{24} \left(\frac{x^2}{2} + \frac{1}{2} - \frac{1}{2} x^3 \right) \Big|_0^1 \\
 &= 1 - 0.02083 = 0.9792
 \end{aligned}$$

d) $P(X < 2, Y < 2)$ is the integral of $f_{XY}(x, y)$ over the following region.



$$\begin{aligned}
 E(X) &= \frac{1}{24} \int_0^3 \int_x^{x+2} x(x+y) dy dx = \frac{1}{24} \int_0^3 x^2 y + \frac{xy^2}{2} \Big|_x^{x+2} dx \\
 &= \frac{1}{24} \int_0^3 (4x^2 + 2x) dx = \frac{1}{24} \left[\frac{4x^3}{3} + x^2 \Big|_0^3 \right] = \frac{15}{8}
 \end{aligned}$$

e)

$$\begin{aligned}
 E(X) &= \frac{1}{24} \int_0^3 \int_x^{x+2} x(x+y) dy dx = \frac{1}{24} \int_0^3 x^2 y + \frac{xy^2}{2} \Big|_x^{x+2} dx \\
 &= \frac{1}{24} \int_0^3 (4x^2 + 2x) dx = \frac{1}{24} \left[\frac{4x^3}{3} + x^2 \Big|_0^3 \right] = \frac{15}{8}
 \end{aligned}$$

f)

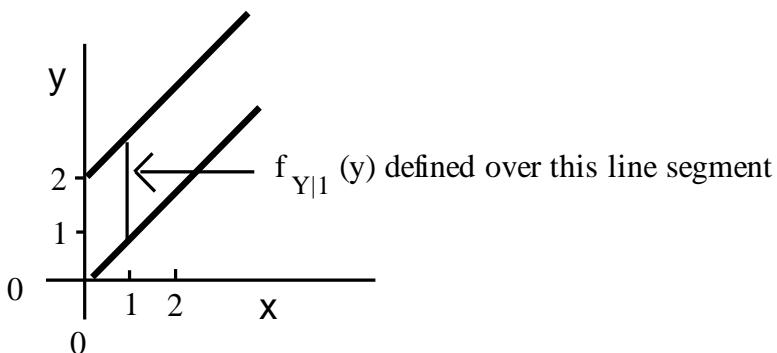
$$\begin{aligned}
 V(X) &= \frac{1}{24} \int_0^3 \int_x^{x+2} x^2(x+y) dy dx - \left(\frac{15}{8} \right)^2 = \frac{1}{24} \int_0^3 x^3 y + \frac{x^2 y^2}{2} \Big|_x^{x+2} dx - \left(\frac{15}{8} \right)^2 \\
 &= \frac{1}{24} \int_0^3 (3x^3 + 4x^2 + 4x - \frac{x^4}{4}) dx - \left(\frac{15}{8} \right)^2 \\
 &= \frac{1}{24} \left[\frac{3x^4}{4} + \frac{4x^3}{3} + 2x^2 - \frac{x^5}{20} \Big|_0^3 \right] - \left(\frac{15}{8} \right)^2 = \frac{31707}{320}
 \end{aligned}$$

g) $f_X(x)$ is the integral of $f_{XY}(x, y)$ over the interval from x to $x+2$. That is,

$$f_X(x) = \frac{1}{24} \int_x^{x+2} (x+y) dy = \frac{1}{24} \left[xy + \frac{y^2}{2} \Big|_x^{x+2} \right] = \frac{x}{6} + \frac{1}{12} \text{ for } 0 < x < 3.$$

h) $f_{Y|1}(y) = \frac{f_{XY}(1,y)}{f_X(1)} = \frac{\frac{1}{24}(1+y)}{\frac{1}{6} + \frac{1}{12}} = \frac{1+y}{6}$ for $1 < y < 3$.

See the following graph,

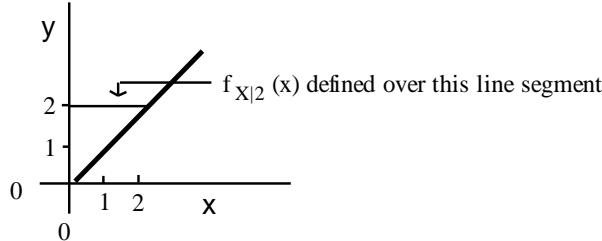


i) $E(Y|X=1) = \int_1^3 y \left(\frac{1+y}{6} \right) dy = \frac{1}{6} \int_1^3 (y + y^2) dy = \frac{1}{6} \left(\frac{y^2}{2} + \frac{y^3}{3} \right) \Big|_1^3 = 2.111$

j) $P(Y > 2 | X = 1) = \int_2^3 \left(\frac{1+y}{6} \right) dy = \frac{1}{6} \int_2^3 (1+y) dy = \frac{1}{6} \left[y + \frac{y^2}{2} \right]_2^3 = 0.5833$

k) $f_{X|2}(x) = \frac{f_{XY}(x, 2)}{f_Y(2)}$. Here $f_Y(y)$ is determined by integrating over x . There are three regions of integration. For $0 < y \leq 2$ the integration is from 0 to y . For $2 < y \leq 3$ the integration is from $y-2$ to y . For $3 < y < 5$ the integration is from y to 3. Because the condition is $y=2$, only the first integration is needed. $f_Y(y) = \frac{1}{24} \int_0^y (x+y) dx = \frac{1}{24} \left[\frac{x^2}{2} + xy \right]_0^y = \frac{y^2}{16}$

for $0 < y \leq 2$.



Therefore, $f_Y(2) = 1/4$ and $f_{X|2}(x) = \frac{\frac{1}{24}(x+2)}{1/4} = \frac{x+2}{6}$ for $0 < x < 3$

5-15. $c \int_0^3 \int_0^x xy dy dx = c \int_0^3 x \frac{y^2}{2} \Big|_0^x dx = c \int_0^3 \frac{x^3}{2} dx \frac{x^4}{8} = \frac{81}{8} c$. Therefore, $c = 8/81$

a) $P(X < 1, Y < 2) = \frac{8}{81} \int_0^1 \int_0^x xy dy dx = \frac{8}{81} \int_0^1 x \frac{x^3}{2} dx = \frac{8}{81} \left(\frac{1}{8} \right) = \frac{1}{81}$.

b) $P(1 < X < 2) = \frac{8}{81} \int_1^2 \int_0^x xy dy dx = \frac{8}{81} \int_1^2 x \frac{x^2}{2} dx = \left(\frac{8}{81} \right) \frac{x^4}{8} \Big|_1^2 = \left(\frac{8}{81} \right) \frac{(2^4 - 1)}{8} = \frac{5}{27}$.

c)

$$\begin{aligned} P(Y > 1) &= \frac{8}{81} \int_1^3 \int_0^x xy dy dx = \frac{8}{81} \int_1^3 x \left(\frac{x^2 - 1}{2} \right) dx = \frac{8}{81} \int_1^3 \frac{x^3}{2} - \frac{x}{2} dx = \frac{8}{81} \left(\frac{x^4}{8} - \frac{x^2}{4} \right) \Big|_1^3 \\ &= \frac{8}{81} \left[\left(\frac{3^4}{8} - \frac{3^2}{4} \right) - \left(\frac{1^4}{8} - \frac{1^2}{4} \right) \right] = \frac{64}{81} = 0.7901 \end{aligned}$$

d) $P(X < 2, Y < 2) = \frac{8}{81} \int_0^2 \int_0^x xy dy dx = \frac{8}{81} \int_0^2 x \frac{x^3}{2} dx = \frac{8}{81} \left(\frac{2^4}{8} \right) = \frac{16}{81}$.

e)

$$\begin{aligned}
 E(X) &= \frac{8}{81} \int_0^3 \int_0^x xy dy dx = \frac{8}{81} \int_0^3 \int_0^x x^2 y dy dx = \frac{8}{81} \int_0^3 \frac{x^2}{2} x^2 dx = \frac{8}{81} \int_0^3 \frac{x^4}{2} dx \\
 &= \left(\frac{8}{81} \right) \left(\frac{3^5}{10} \right) = \frac{12}{5}
 \end{aligned}$$

f)

$$\begin{aligned}
 E(Y) &= \frac{8}{81} \int_0^3 \int_0^x y(xy) dy dx = \frac{8}{81} \int_0^3 \int_0^x xy^2 dy dx = \frac{8}{81} \int_0^3 x \frac{x^3}{3} dx \\
 &= \frac{8}{81} \int_0^3 \frac{x^4}{3} dx = \left(\frac{8}{81} \right) \left(\frac{3^5}{15} \right) = \frac{8}{5}
 \end{aligned}$$

g) $f(x) = \frac{8}{81} \int_0^x y dy = \frac{4x^3}{81} \quad 0 < x < 3$

h) $f_{Y|x=1}(y) = \frac{f(1, y)}{f(1)} = \frac{\frac{8}{81}(1)y}{\frac{4(1)^3}{81}} = 2y \quad 0 < y < 1$

i) $E(Y | X = 1) = \int_0^1 2y dy = y^2 \Big|_0^1 = 1$

j) $P(Y > 2 | X = 1) = 0$ this isn't possible since the values of y are $0 < y < x$.

k) $f(y) = \frac{8}{81} \int_0^3 xy dx = \frac{4y}{9}$, therefore

$$f_{X|Y=2}(x) = \frac{f(x, 2)}{f(2)} = \frac{\frac{8}{81}x(2)}{\frac{4(2)}{9}} = \frac{2x}{9} \quad 0 < x < 3$$

5-16. Solve for c

$$\begin{aligned}
 c \int_0^\infty \int_0^x e^{-3x-4y} dy dx &= \frac{c}{4} \int_0^\infty e^{-3x} (1 - e^{-4x}) dx = \frac{c}{4} \int_0^\infty e^{-3x} - e^{-7x} dx = \\
 \frac{c}{4} \left(\frac{1}{3} - \frac{1}{7} \right) &= \frac{1}{21}c. \quad c = 21
 \end{aligned}$$

$$\begin{aligned}
 \text{a) } P(X < 1, Y < 2) &= 21 \int_0^1 \int_0^x e^{-3x-4y} dy dx = \frac{21}{4} \int_0^1 e^{-3x} (1 - e^{-4x}) dx = \frac{21}{4} \int_0^1 e^{-3x} - e^{-7x} dx \\
 &= \frac{21}{4} \left(\frac{e^{-7x}}{7} - \frac{e^{-3x}}{3} \right) \Big|_0^1 = 0.9135
 \end{aligned}$$

$$\text{b) } P(1 < X < 2) = 21 \int_1^2 \int_0^x e^{-3x-4y} dy dx = \frac{21}{4} \int_1^2 (e^{-3x} - e^{-7x}) dx$$

$$= \frac{21}{4} \left(\frac{e^{-7x}}{7} - \frac{e^{-3x}}{3} \right) \Big|_1^2 = 0.1821$$

$$\text{c) } P(Y > 3) = 21 \int_{3}^{\infty} \int_{3}^{x} e^{-3x-4y} dy dx = \frac{21}{4} \int_{3}^{\infty} e^{-3x} (e^{-12} - e^{-4x}) dx \\ = \frac{21}{4} \left(\frac{e^{-7x}}{7} - \frac{e^{-12} e^{-3x}}{3} \right) \Big|_3^{\infty} = 7.583 \times 10^{-10}$$

$$\text{d) } P(X < 2, Y < 2) = 21 \int_{0}^2 \int_{0}^x e^{-3x-4y} dy dx = \frac{21}{4} \int_{0}^2 e^{-3x} (1 - e^{-4x}) dx = \frac{21}{4} \left[\left(\frac{e^{-14}}{7} - \frac{e^{-6}}{3} \right) - \left(\frac{e^0}{7} - \frac{e^0}{3} \right) \right]_0^2 \\ = 0.9957$$

$$\text{e) } E(X) = 21 \int_{0}^{\infty} \int_{0}^x x e^{-3x-4y} dy dx = \frac{10}{21}$$

$$\text{f) } E(Y) = 21 \int_{0}^{\infty} \int_{0}^x y e^{-3x-4y} dy dx = \frac{29}{56}$$

$$\text{g) } f(x) = 21 \int_{0}^x e^{-3x-4y} dy = \frac{21 e^{-3x}}{4} (1 - e^{-4x}) = \frac{21}{4} (e^{-3x} - e^{-7x}) \text{ for } 0 < x$$

$$\text{h) } f_{Y|X=1}(y) = \frac{f_{X,Y}(1,y)}{f_X(1)} = \frac{21 e^{-3-4y}}{\frac{21}{4} (e^{-3} - e^{-7})} = 4.075 e^{-4y} \quad 0 < y < 1$$

$$\text{i) } E(Y/X=1) = 4.075 \int_0^1 y e^{-4y} dy = 0.2314$$

$$\text{j) } f_{X|Y=2}(x) = \frac{f_{X,Y}(x,2)}{f_Y(2)} = \frac{21 e^{-3x-8}}{7 e^{-14}} = 3 e^{-3x+6} \text{ for } 2 < x, \\ \text{where } f(y) = 7 e^{-7y} \text{ for } 0 < y$$

$$5-17. \quad c \int_{0}^{\infty} \int_{x}^{\infty} e^{-2x} e^{-3y} dy dx = \frac{c}{3} \int_{0}^{\infty} e^{-2x} (e^{-3x}) dx = \frac{c}{3} \int_{0}^{\infty} e^{-5x} dx = \frac{1}{15} c \quad c = 15$$

a)

$$P(X < 1, Y < 2) = 15 \int_{0}^1 \int_{x}^2 e^{-2x-3y} dy dx = 5 \int_{0}^1 e^{-2x} (e^{-3x} - e^{-6}) dx \\ = 5 \int_{0}^1 e^{-5x} dx - 5 e^{-6} \int_{0}^1 e^{-2x} dx = 1 - e^{-5} + \frac{5}{2} e^{-6} (e^{-2} - 1) = 0.9879$$

$$\text{b) } P(1 < X < 2) = 15 \int_{1}^2 \int_{x}^{\infty} e^{-2x-3y} dy dx = 5 \int_{1}^2 e^{-5x} dy dx = (e^{-5} - e^{-10}) = 0.0067$$

c)

$$\begin{aligned}
 P(Y > 3) &= 15 \left(\int_0^3 \int_3^\infty e^{-2x-3y} dy dx + \int_3^\infty \int_x^\infty e^{-2x-3y} dy dx \right) = 5 \int_0^3 e^{-9} e^{-2x} dx + 5 \int_3^\infty e^{-5x} dx \\
 &= -\frac{3}{2} e^{-15} + \frac{5}{2} e^{-9} = 0.000308
 \end{aligned}$$

d)

$$\begin{aligned}
 P(X < 2, Y < 2) &= 15 \int_0^2 \int_0^2 e^{-2x-3y} dy dx = 5 \int_0^2 e^{-2x} (e^{-3x} - e^{-6}) dx = \\
 &= 5 \int_0^2 e^{-5x} dx - 5e^{-6} \int_0^2 e^{-2x} dx = (1 - e^{-10}) + \frac{5}{2} e^{-6} (e^{-4} - 1) = 0.9939
 \end{aligned}$$

$$\text{e) } E(X) = 15 \int_0^\infty \int_x^\infty x e^{-2x-3y} dy dx = 5 \int_0^\infty x e^{-5x} dx = \frac{1}{5^2} = 0.04$$

f)

$$\begin{aligned}
 E(Y) &= 15 \int_0^\infty \int_x^\infty y e^{-2x-3y} dy dx = \frac{-3}{2} \int_0^\infty 5ye^{-5y} dy + \frac{5}{2} \int_0^\infty 3ye^{-3y} dy \\
 &= -\frac{3}{10} + \frac{5}{6} = \frac{8}{15}
 \end{aligned}$$

$$\text{g) } f(x) = 15 \int_x^\infty e^{-2x-3y} dy = \frac{15}{3} (e^{-2z-3x}) = 5e^{-5x} \text{ for } x > 0$$

$$\text{h) } f_X(1) = 5e^{-5} \quad f_{XY}(1, y) = 15e^{-2-3y}$$

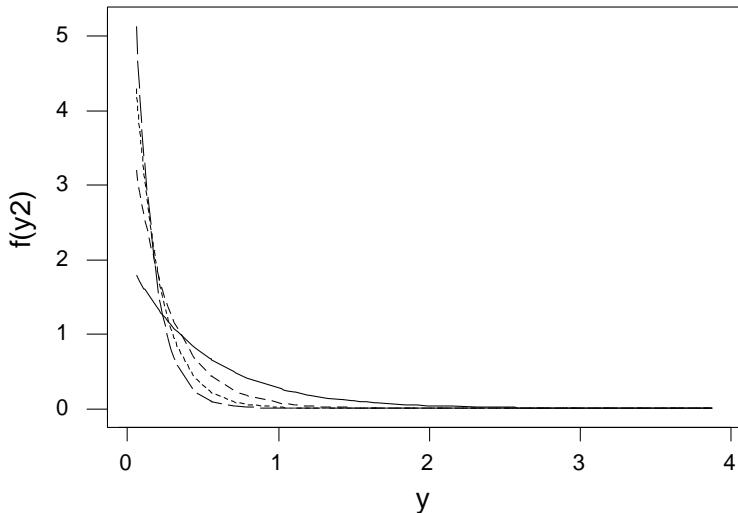
$$f_{Y|X=1}(y) = \frac{15e^{-2-3y}}{5e^{-5}} = 3e^{3-3y} \text{ for } 1 < y$$

$$\text{i) } E(Y | X = 1) = \int_1^\infty 3ye^{3-3y} dy = -ye^{3-3y} \Big|_1^\infty + \int_1^\infty e^{3-3y} dy = 4/3$$

$$\text{j) } \int_1^2 3e^{3-3y} dy = 1 - e^{-3} = 0.9502 \text{ for } 0 < y, \quad f_Y(2) = \frac{15}{2} e^{-6}$$

$$\text{k) For } y > 0 \quad f_{X|Y=1}(y) = \frac{15e^{-2x-3}}{\frac{15}{2}e^{-3}} = 2e^{-2x} \text{ for } 0 < x < 1$$

5-18. a) $f_{Y|X=x}(y)$, for $x = 2, 4, 6, 8$



b) $P(Y < 2 | X = 2) = \int_0^2 2e^{-2y} dy = 0.9817$

c) $E(Y | X = 2) = \int_0^\infty 2ye^{-2y} dy = 1/2$ (using integration by parts)

d) $E(Y | X = x) = \int_0^\infty xye^{-xy} dy = 1/x$ (using integration by parts)

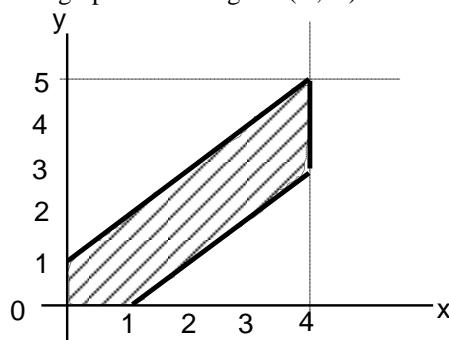
e) Use $f_X(x) = \frac{1}{b-a} = \frac{1}{10}$, $f_{Y|X}(x, y) = xe^{-xy}$, and the relationship

$$f_{Y|X}(x, y) = \frac{f_{XY}(x, y)}{f_X(x)}$$

Therefore, $xe^{-xy} = \frac{f_{XY}(x, y)}{1/10}$ and $f_{XY}(x, y) = \frac{xe^{-xy}}{10}$

f) $f_Y(y) = \int_0^{10} \frac{xe^{-xy}}{10} dx = \frac{1 - 10ye^{-10y} - e^{-10y}}{10y^2}$ (using integration by parts)

- 5-19. The graph of the range of (X, Y) is



$$\begin{aligned} & \int_0^1 \int_0^{x+1} c dy dx + \int_1^4 \int_{x-1}^{x+1} c dy dx = 1 \\ &= c \int_0^1 (x+1) dx + 2c \int_1^4 dx \\ &= \frac{3}{2}c + 6c = 7.5c = 1 \end{aligned}$$

Therefore, $c = 1/7.5 = 2/15$

a) $P(X < 0.5, Y < 1) = \int_0^{0.5} \int_0^1 \frac{1}{7.5} dy dx = \frac{1}{15}$

b) $P(X < 0.5) = \int_0^{0.5} \int_0^{x+1} \frac{1}{7.5} dy dx = \frac{1}{7.5} \int_0^{0.5} (x+1) dx = \frac{2}{15} \left(\frac{5}{8}\right) = \frac{1}{12}$

c)

$$\begin{aligned} E(X) &= \int_0^1 \int_0^{x+1} \frac{x}{7.5} dy dx + \int_1^4 \int_{x-1}^{x+1} \frac{x}{7.5} dy dx \\ &= \frac{1}{7.5} \int_0^1 (x^2 + x) dx + \frac{2}{7.5} \int_1^4 (x) dx = \frac{12}{15} \left(\frac{5}{6}\right) + \frac{2}{7.5} (7.5) = \frac{19}{9} \end{aligned}$$

d)

$$\begin{aligned} E(Y) &= \frac{1}{7.5} \int_0^1 \int_0^{x+1} y dy dx + \frac{1}{7.5} \int_1^4 \int_{x-1}^{x+1} y dy dx \\ &= \frac{1}{7.5} \int_0^1 \frac{(x+1)^2}{2} dx + \frac{1}{7.5} \int_1^4 \frac{(x+1)^2 - (x-1)^2}{2} dx \\ &= \frac{1}{15} \int_0^1 (x^2 + 2x + 1) dx + \frac{1}{15} \int_1^4 4x dx \\ &= \frac{1}{15} \left(\frac{7}{3}\right) + \frac{1}{15} (30) = \frac{97}{45} \end{aligned}$$

e)

$$f(x) = \int_0^{x+1} \frac{1}{7.5} dy = \left(\frac{x+1}{7.5} \right) \quad \text{for } 0 < x < 1,$$

$$f(x) = \int_{x-1}^{x+1} \frac{1}{7.5} dy = \left(\frac{x+1-(x-1)}{7.5} \right) = \frac{2}{7.5} \quad \text{for } 1 < x < 4$$

f)

$$f_{Y|X=1}(y) = \frac{f_{XY}(1, y)}{f_X(1)} = \frac{1/7.5}{2/7.5} = 0.5$$

$$f_{Y|X=1}(y) = 0.5 \quad \text{for } 0 < y < 2$$

g) $E(Y | X = 1) = \int_0^2 \frac{y}{2} dy = \left. \frac{y^2}{4} \right|_0^2 = 1$

$$\text{h) } P(Y < 0.5 | X = 1) = \int_0^{0.5} 0.5 dy = 0.5y \Big|_0^{0.5} = 0.25$$

- 5-20. Let X, Y, and Z denote the time until a problem on line 1, 2, and 3, respectively.

a)

$$P(X > 40, Y > 40, Z > 40) = [P(X > 40)]^3$$

because the random variables are independent with the same distribution. Now,

$$P(X > 40) = \int_{40}^{\infty} \frac{1}{40} e^{-x/40} dx = -e^{-x/40} \Big|_{40}^{\infty} = e^{-1} \text{ and the answer is}$$

$$(e^{-1})^3 = e^{-3} = 0.0498$$

$$\text{b) } P(30 < X < 40, 30 < Y < 40, 30 < Z < 40) = [P(30 < X < 40)]^3 \text{ and}$$

$$P(30 < X < 40) = -e^{-x/40} \Big|_{30}^{40} = e^{-0.75} - e^{-1} = 0.1045.$$

The answer is $0.1045^3 = 0.0011$.

c) The joint density is not needed because the process is represented by three independent exponential distributions. Therefore, the probabilities may be multiplied.

- 5-21. $\mu = 3.2, \lambda = 1/3.2$

$$\begin{aligned} P(X > 5, Y > 5) &= (1/10.24) \int_5^{\infty} \int_5^{\infty} e^{-\frac{x}{3.2} - \frac{y}{3.2}} dy dx = 3.2 \int_5^{\infty} e^{-\frac{x}{3.2}} \left(e^{-\frac{5}{3.2}} \right) dx \\ &= \left(e^{-\frac{5}{3.2}} \right) \left(e^{-\frac{5}{3.2}} \right) = 0.0439 \end{aligned}$$

$$\begin{aligned} P(X > 10, Y > 10) &= (1/10.24) \int_{10}^{\infty} \int_{10}^{\infty} e^{-\frac{x}{3.2} - \frac{y}{3.2}} dy dx = 3.2 \int_{10}^{\infty} e^{-\frac{x}{3.2}} \left(e^{-\frac{10}{3.2}} \right) dx \\ &= \left(e^{-\frac{10}{3.2}} \right) \left(e^{-\frac{10}{3.2}} \right) = 0.0019 \end{aligned}$$

b) Let X denote the number of orders in a 5-minute interval. Then X is a Poisson random variable with $\lambda = 5/3.2 = 1.5625$.

$$P(X = 1) = \frac{e^{-1.5625}(1.5625)^1}{1!} = 0.3275$$

For both systems, $P(X = 1)P(Y = 1) = 0.3275^2 = 0.1073$

c) The joint probability distribution is not necessary because the two processes are independent and we can just multiply the probabilities.

- 5-22. (a) X: the life time of blade and Y: the life time of bearing
 $f(x) = (1/3)e^{-x/3}$ $f(y) = (1/4)e^{-y/4}$

$$P(X > 6, Y > 6) = P(X > 6)P(Y > 6) = e^{-6/3}e^{-6/4} = 0.0302$$

(b) $P(X > t, Y > t) = e^{-t^3}e^{-t^4} = e^{-7t/12} = 0.95 \rightarrow t = -12 \ln(0.95)/7 = 0.0879$ years

5-23. a) $P(X < 0.5) = \int_0^{0.5} \int_0^1 \int_0^1 (10xyz) dz dy dx = \int_0^{0.5} \int_0^1 (5xy) dy dx = \int_0^{0.5} (2.5x) dx = 1.25x^2 \Big|_0^{0.5} = 0.3125$

b) $P(X < 0.5, Y < 0.5) = \int_0^{0.5} \int_0^{0.5} \int_0^1 (10xyz) dz dy dx$
 $= \int_0^{0.5} \int_0^{0.5} (5xy) dy dx = \int_0^{0.5} (0.625x) dx = \frac{0.625x^2}{2} \Big|_0^{0.5} = 0.0781$

c) $P(Z < 2) = 1$, because the range of Z is from 0 to 1.

d) $P(X < 0.5 \text{ or } Z < 2) = P(X < 0.5) + P(Z < 2) - P(X < 0.5, Z < 2)$. Now, $P(Z < 2) = 1$ and $P(X < 0.5, Z < 2) = P(X < 0.5)$. Therefore, the answer is 1.

e) $E(X) = \int_0^1 \int_0^1 \int_0^1 (10x^2yz) dz dy dx = \int_0^1 (2.5x^2) dx = \frac{2.5x^3}{3} \Big|_0^1 = 0.833$

f) $P(X < 0.5 | Y = 0.5)$ is the integral of the conditional density $f_{X|Y}(x)$. Now,

$$f_{X|0.5}(x) = \frac{f_{XY}(x, 0.5)}{f_Y(0.5)} \text{ and } f_{XY}(x, 0.5) = \int_0^1 (10x(0.5)z) dz = 5x0.5 = 2.5x \text{ for } 0 < x < 1$$

and $0 < y < 1$. Also, $f_Y(y) = \int_0^1 \int_0^1 (10xyz) dz dx = 2.5y$ for $0 < y < 1$; $f_Y(0.5) = 1.25$

Therefore, $f_{X|0.5}(x) = \frac{2.5x}{1.25} = 2x$ for $0 < x < 1$.

Then, $P(X < 0.5 | Y = 0.5) = \int_0^{0.5} 2x dx = 0.25$.

g) $P(X < 0.5, Y < 0.5 | Z = 0.8)$ is the integral of the conditional density of X and Y. Now,

$f_Z(z) = 2.5z$ for $0 < z < 1$ as in part a) and

$$f_{XY|Z}(x, y) = \frac{f_{XYZ}(x, y, z)}{f_Z(z)} = \frac{10xy(0.8)}{2.5(0.8)} = 4xy \text{ for } 0 < x < 1 \text{ and } 0 < y < 1.$$

Then, $P(X < 0.5, Y < 0.5 | Z = 0.8) = \int_0^{0.5} \int_0^{0.5} (4xy) dy dx = \int_0^{0.5} (x/2) dx = \frac{1}{16} = 0.0625$

h) $f_{YZ}(y, z) = \int_0^1 (10xyz) dx = 5yz$ for $0 < y < 1$ and $0 < z < 1$.

Then, $f_{X|YZ}(x) = \frac{f_{XYZ}(x, y, z)}{f_{YZ}(y, z)} = \frac{10x(0.5)(0.8)}{5(0.5)(0.8)} = 2x$ for $0 < x < 1$.

i) Therefore, $P(X < 0.5 | Y = 0.5, Z = 0.8) = \int_0^{0.5} 2x dx = 0.25$

5-24. $\iint \int_0^4 c dz dy dx$ = the volume of a cylinder with a base of radius 2 and a height of 4 =

$$(\pi 2^2)4 = 16\pi. \text{ Therefore, } c = \frac{1}{16\pi}$$

a) $P(X^2 + Y^2 < 2)$ equals the volume of a cylinder of radius $\sqrt{2}$ and a height of 4 ($= 8\pi$) times c . Therefore, the answer is $\frac{8\pi}{16\pi} = 1/2$.

b) $P(Z < 2)$ equals half the volume of the region where $f_{XYZ}(x, y, z)$ is positive times 1/c. Therefore, the answer is 0.5.

$$c) E(X) = \int_{-2}^2 \int_{-\sqrt{4-x^2}}^{\sqrt{4-x^2}} \int_0^4 \frac{x}{c} dz dy dx = c \int_{-2}^2 \left[4xy \right]_{-\sqrt{4-x^2}}^{\sqrt{4-x^2}} dx = c \int_{-2}^2 (8x\sqrt{4-x^2}) dx.$$

Using substitution, $u = 4 - x^2$, $du = -2x dx$, and

$$E(X) = c \int 4\sqrt{u} du = \frac{-4}{c} \frac{2}{3} (4 - x^2)^{\frac{3}{2}} \Big|_{-2}^2 = 0$$

$$d) f_{X|1}(x) = \frac{f_{XY}(x, 1)}{f_Y(1)} \text{ and } f_{XY}(x, y) = c \int_0^4 dz = \frac{4}{c} = \frac{1}{4\pi} \text{ for } x^2 + y^2 < 4.$$

$$\text{Also, } f_Y(y) = c \int_{-\sqrt{4-y^2}}^{\sqrt{4-y^2}} \int_0^4 dz dx = 8c\sqrt{4-y^2} \text{ for } -2 < y < 2.$$

$$\text{Then, } f_{X|y}(x) = \frac{4c}{8c\sqrt{4-y^2}} \text{ evaluated at } y = 1. \text{ That is, } f_{X|1}(x) = \frac{1}{2\sqrt{3}} \text{ for } -\sqrt{3} < x < \sqrt{3}$$

$$\text{Therefore, } P(X < 1 | Y < 1) = \int_{-\sqrt{3}}^1 \frac{1}{2\sqrt{3}} dx = \frac{1 + \sqrt{3}}{2\sqrt{3}} = 0.7887$$

$$e) f_{XY|1}(x, y) = \frac{f_{XYZ}(x, y, 1)}{f_Z(1)} \text{ and } f_Z(z) = \int_{-2}^2 \int_{-\sqrt{4-x^2}}^{\sqrt{4-x^2}} c dy dx = \int_{-2}^2 2c\sqrt{4-x^2} dx$$

Because $f_Z(z)$ is a density over the range $0 < z < 4$ that does not depend on Z, $f_Z(z) = 1/4$ for $0 < z < 4$. Then, $f_{XY|1}(x, y) = \frac{c}{1/4} = \frac{1}{4\pi}$ for $x^2 + y^2 < 4$.

$$\text{Then, } P(X^2 + Y^2 < 1 | Z = 1) = \frac{\text{area in } x^2 + y^2 < 1}{4\pi} = 1/4$$

$$f) f_{Z|xy}(z) = \frac{f_{XYZ}(x, y, z)}{f_{XY}(x, y)} \text{ and } f_{XY}(x, y) = \frac{1}{4\pi} \text{ for } x^2 + y^2 < 4. \text{ Therefore,}$$

$$f_{Z|xy}(z) = \frac{\frac{1}{16\pi}}{\frac{1}{4\pi}} = 1/4 \text{ for } 0 < z < 4.$$

5-25. Determine c such that $f(xyz) = c$ is a joint density probability over the region $x > 0$, $y > 0$ and $z > 0$ with $x + y + z < 1$

$$\begin{aligned}
 f(xyz) &= c \int_0^1 \int_0^{1-x} \int_0^{1-x-y} dz dy dx = \int_0^1 \int_0^{1-x} c(1-x-y) dy dx = \int_0^1 \left(c(y - xy - \frac{y^2}{2}) \Big|_0^{1-x} \right) dx \\
 &= \int_0^1 c \left((1-x) - x(1-x) - \frac{(1-x)^2}{2} \right) dx = \int_0^1 c \left(\frac{(1-x)^2}{2} \right) dx = c \left(\frac{1}{2}x - \frac{x^2}{2} + \frac{x^3}{6} \right) \Big|_0^1 \\
 &= c \frac{1}{6}. \quad \text{Therefore, } c = 6.
 \end{aligned}$$

a) $P(X < 0.25, Y < 0.25, Z < 0.25) = 6 \int_0^{1-x} \int_0^{1-x-y} dz dy dx \Rightarrow$ The conditions $x < 0.25, y < 0.25,$

$z < 0.25$ and $x+y+z < 1$ make a space that is a cube with a volume of 0.015625. Therefore the probability of $P(X < 0.25, Y < 0.25, Z < 0.25) = 6(0.015625) = 0.09375$

b)

$$\begin{aligned}
 P(X < 0.5, Y < 0.5) &= \int_0^{0.5} \int_0^{0.5} 6(1-x-y) dy dx = \int_0^{0.5} \left(6y - 6xy - 3y^2 \right) \Big|_0^{0.5} dx \\
 &= \int_0^{0.5} \left(\frac{9}{4} - 3x \right) dx = \left(\frac{9}{4}x - \frac{3}{2}x^2 \right) \Big|_0^{0.5} = 3/4
 \end{aligned}$$

c)

$$\begin{aligned}
 P(X < 0.5) &= 6 \int_0^{0.5} \int_0^{1-x} \int_0^{1-x-y} dz dy dx = \int_0^{0.5} \int_0^{1-x} 6(1-x-y) dy dx = \int_0^{0.5} 6\left(y - xy - \frac{y^2}{2}\right) \Big|_0^{1-x} dy \\
 &= \int_0^{0.5} 6\left(\frac{x^2}{2} - x + \frac{1}{2}\right) dx = \left(x^3 - 3x^2 + 3x\right) \Big|_0^{0.5} = 0.875
 \end{aligned}$$

d)

$$\begin{aligned}
 E(X) &= 6 \int_0^1 \int_0^{1-x} \int_0^{1-x-y} x dz dy dx = \int_0^1 \int_0^{1-x} 6x(1-x-y) dy dx = \int_0^1 6x\left(y - xy - \frac{y^2}{2}\right) \Big|_0^{1-x} dy \\
 &= \int_0^1 6\left(\frac{x^3}{2} - x^2 + \frac{x}{2}\right) dx = \left(\frac{3x^4}{4} - 2x^3 + \frac{3x^2}{2}\right) \Big|_0^1 = 0.25
 \end{aligned}$$

e)

$$\begin{aligned}
 f(x) &= 6 \int_0^{1-x} \int_0^{1-x-y} dz dy = \int_0^{1-x} 6(1-x-y) dy = 6 \left(y - xy - \frac{y^2}{2} \right) \Big|_0^{1-x} \\
 &= 6\left(\frac{x^2}{2} - x + \frac{1}{2}\right) = 3(x-1)^2 \text{ for } 0 < x < 1
 \end{aligned}$$

f)

$$f(x, y) = 6 \int_0^{1-x-y} dz = 6(1-x-y)$$

for $x > 0, y > 0$ and $x + y < 1$

g)

$$f(x | y = 0.5, z = 0.5) = \frac{f(x, y = 0.5, z = 0.5)}{f(y = 0.5, z = 0.5)} = \frac{6}{6} = 1 \text{ for } x > 0$$

h) The marginal $f_Y(y)$ is similar to $f_X(x)$ and $f_Y(y) = 3(1-y)^2$ for $0 < y < 1$.

$$f_{X|Y}(x | 0.5) = \frac{f(x, 0.5)}{f_Y(0.5)} = \frac{6(0.5-x)}{3(0.25)} = 4(1-2x) \text{ for } x < 0.5$$

- 5-26. Let X denote the production yield on a day. Then,

$$P(X > 635) = P(Z > \frac{635-680}{45}) = P(Z > -1) = 0.84134.$$

a) Let Y denote the number of days out of five such that the yield exceeds 635. Then, by independence, Y has a binomial distribution with $n = 5$ and $p = 0.84134$. Therefore, the answer is $P(Y = 5) = \binom{5}{5} 0.84134^5 (1-0.84134)^0 = 0.4215$.

b) As in part (a), the answer is

$$P(Y \geq 4) = P(Y = 4) + P(Y = 5)$$

$$= \binom{5}{4} 0.84134^4 (1-0.84134)^1 + 0.4215 = 0.8190$$

- 5-27. a) Let X denote the weight of a brick. Then,

$$P(X > 2.75) = P(Z > \frac{1.2-1.5}{0.3}) = P(Z > -1) = 0.84134.$$

Let Y denote the number of bricks in the sample of 20 that exceed 1.2 kg. Then, by independence, Y has a binomial distribution with $n = 20$ and $p = 0.84134$. Therefore, the answer is

$$P(Y = 20) = \binom{20}{20} 0.84134^{20} = 0.032.$$

b) Let A denote the event that the heaviest brick in the sample exceeds 1.8 kg. Then, $P(A) = 1 - P(A')$ and A' is the event that all bricks weigh less than 1.8 kg. As in part a., $P(X < 1.8) = P(Z < 1)$ and $P(A) = 1 - [P(Z < 1)]^{20} = 1 - 0.84134^{20} = 0.9684$.

- 5-28. a) Let X denote the grams of luminescent ink. Then,

$$P(X < 1.14) = P(Z < \frac{1.14-1.2}{0.3}) = P(Z < -2) = 0.022750$$

Let Y denote the number of bulbs in the sample of 25 that have less than 1.14 grams. Then, by independence, Y has a binomial distribution with $n = 25$ and $p = 0.022750$. Therefore, the answer is $P(Y \geq 1) = 1 - P(Y = 0) = \binom{25}{0} 0.022750^0 (0.97725)^{25} = 1 - 0.5625 = 0.4375$.

b)

$$\begin{aligned} P(Y \leq 5) &= P(Y = 0) + P(Y = 1) + P(Y = 2) + P(Y = 3) + P(Y = 4) + P(Y = 5) \\ &= \binom{25}{0} 0.022750^0 (0.97725)^{25} + \binom{25}{1} 0.022750^1 (0.97725)^{24} + \binom{25}{2} 0.022750^2 (0.97725)^{23} \\ &\quad + \binom{25}{3} 0.022750^3 (0.97725)^{22} + \binom{25}{4} 0.022750^4 (0.97725)^{21} + \binom{25}{5} 0.022750^5 (0.97725)^{20} \\ &= 0.5625 + 0.3274 + 0.09146 + 0.01632 + 0.002090 + 0.0002043 = 0.99997 \approx 1 \end{aligned}$$

$$c) P(Y = 0) = \binom{25}{0} 0.022750^0 (0.97725)^{25} = 0.5625$$

- d) The lamps are normally and independently distributed. Therefore, the probabilities can be multiplied.

Section 5-2

5-29. $E(X) = 2(1/8)+1(1/4)+2(1/2)+4(1/8) = 2$
 $E(Y) = 3(1/8)+4(1/4)+5(1/2)+6(1/8) = 37/8 = 4.625$

$$E(XY) = [2 \times 3 \times (1/8)] + [1 \times 4 \times (1/4)] + [2 \times 5 \times (1/2)] + [4 \times 6 \times (1/8)] \\ = 39/4 = 9.75$$

$$\sigma_{XY} = E(XY) - E(X)E(Y) = 9.75 - (2)(4.625) = 0.5$$

$$V(X) = 2^2(1/8) + 1^2(1/4) + 2^2(1/2) + 4^2(1/8) - 2^2 \\ = 0.5 + 0.25 + 2 + 2 - 2^2 = 4.75 - 4 = 0.75$$

$$V(Y) = 3^2(1/8) + 4^2(1/4) + 5^2(1/2) + 6^2(1/8) - (37/8)^2 = 0.7344$$

$$\rho_{XY} = \frac{\sigma_{XY}}{\sigma_X \sigma_Y} = \frac{0.5}{\sqrt{(0.75)(0.7344)}} = 0.6737$$

5-30. $E(X) = -1(1/8) + (-0.5)(1/5) + 0.5(1/2) + 1(7/40) = 0.2$
 $E(Y) = -2(1/8) + (-1)(1/5) + 1(1/2) + 2(7/40) = 0.4$
 $E(XY) = [-1 \times -2 \times (1/8)] + [-0.5 \times -1 \times (1/5)] + [0.5 \times 1 \times (1/2)] + [1 \times 2 \times (7/40)] = 0.95$
 $V(X) = 0.435$
 $V(Y) = 1.74$
 $\sigma_{XY} = 0.95 - (0.2)(0.4) = 0.87$
 $\rho_{XY} = \frac{\sigma_{XY}}{\sigma_X \sigma_Y} = \frac{0.87}{\sqrt{0.435} \sqrt{1.74}} = 1$

5-31.

$$\sum_{x=1}^3 \sum_{y=1}^2 c(x+y) = 21c, \quad c = 1/21$$

$$E(X) = \frac{46}{21} \quad E(Y) = \frac{11}{7} \quad E(XY) = \frac{24}{7} \quad \sigma_{xy} = \frac{24}{7} - \left(\frac{46}{21} \right) \left(\frac{11}{7} \right) = \frac{-2}{147} = -0.0136$$

$$E(X^2) = \frac{114}{21} \quad E(Y^2) = \frac{57}{21}$$

$$V(X) = EX^2 - (EX)^2 = 0.63 \quad V(Y) = 0.24$$

$$\rho = \frac{-0.0136}{\sqrt{0.63} \sqrt{0.24}} = -0.035$$

- 5-32. The marginal distribution of X is

x	f(x)
0	0.75
1	0.2
2	0.05

$$E(X) = 0(0.75) + 1(0.2) + 2(0.05) = 0.3$$

$$E(Y) = 0(0.3) + 1(0.28) + 2(0.25) + 3(0.17) = 1.29$$

$$E(X^2) = 0(0.75) + 1(0.2) + 4(0.05) = 0.4$$

$$E(Y^2) = 0(0.3) + 1(0.28) + 4(0.25) + 9(0.17) = 1.146$$

$$V(X) = 0.4 - 0.3^2 = 0.31$$

$$V(Y) = 2.81 - 1.146^2 = 1.16$$

$$E(XY) = [0 \times 0 \times (0.225)] + [0 \times 1 \times (0.21)] + [0 \times 2 \times (0.1875)] + \dots + [2 \times 3 \times (0.0085)] = 0.387$$

$$\sigma_{XY} = 0.387 - (0.3)(1.29) = 0$$

$$\rho_{XY} = \frac{\sigma_{XY}}{\sigma_X \sigma_Y} = 0$$

- 5-33. Let X and Y denote the number of patients who improve or degrade, respectively, and let Z denote the number of patients that remain the same. If $X = 0$, then Y can equal 0,1,2,3, or 4. However, if $X = 4$ then $Y = 0$. Consequently, the range of the joint distribution of X and Y is not rectangular. Therefore, X and Y are not independent.

$$\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X, Y).$$

Therefore,

$$\text{Cov}(X, Y) = 0.5[\text{Var}(X + Y) - \text{Var}(X) - \text{Var}(Y)]$$

Here X and Y are binomially distributed when considered individually. Therefore,

$$f_X(x) = \frac{4!}{x!(4-x)!} 0.4^x (1-0.4)^{4-x}$$

$$f_Y(y) = \frac{4!}{y!(4-y)!} 0.1^y (1-0.1)^{4-y}$$

And

$$\text{Var}(X) = 4(0.4)(0.6) = 0.96$$

$$\text{Var}(Y) = 4(0.1)(0.9) = 0.36$$

Also, $W = X + Y$ is binomial with $n = 4$, and $p = 0.4 + 0.1 = 0.5$. Therefore,

$$\text{Var}(X + Y) = 4(0.5)(0.5) = 1$$

Therefore, $\text{Cov}(X, Y) = 0.5[1 - 0.96 - 0.36] = -0.16$

$$\rho_{XY} = \frac{\text{Cov}(X, Y)}{\sqrt{V(X)V(Y)}} = \frac{-0.16}{\sqrt{0.96 \times 0.36}} = -0.272$$

- 5-34.

Transaction	Frequency	Selects(X)	Updates(Y)	Inserts(Z)
New Order	43	23	11	12
Payment	44	4.2	3	1
Order Status	4	11.4	0	0
Delivery	5	130	120	0
Stock Level	4	0	0	0
Mean Value		18.694	12.05	5.6

$$(a) \text{ COV}(X, Y) = E(XY) - E(X)E(Y) = 23*11*0.43 + 4.2*3*0.44 + 11.4*0*0.04 + 130*120*0.05 + 0*0*0.04 - 18.694*12.05 = 669.0713$$

$$(b) V(X) = 735.9644, V(Y) = 630.7875; \text{Corr}(X, Y) = \text{cov}(X, Y) / (\sqrt{V(X)V(Y)})^{0.5} = 0.9820$$

- (c) $\text{COV}(Y,Z)=11*12*0.43+3*1*0.44+0-12.05*5.6 = -9.4$
 (d) $V(Z)=31$; $\text{Corr}(X,Z)=-0.067$

5-35. $\int_0^2 \int_0^x cxy dy dx = c \int_0^2 \frac{1}{2} x^3 dx = (\frac{c}{8})(2^4) = 1$, $c = 1/2$, $E(X) = 8/5$, and $E(Y) = 16/15$

$$E(XY) = \frac{1}{2} \int_0^2 \int_0^x xy(xy) dy dx = \frac{16}{9}$$

$$\sigma_{xy} = \frac{16}{9} - \left(\frac{8}{5} \right) \left(\frac{16}{15} \right) = 0.071$$

$$E(X^2) = \frac{8}{3} \quad E(Y^2) = \frac{4}{3}$$

$$V(x) = 0.107, \quad V(Y) = 0.196$$

$$\rho = \frac{0.071}{\sqrt{0.107} \sqrt{0.196}} = 0.492$$

5-36. $\int_0^1 \int_0^{x+1} c dy dx + \int_1^4 \int_{x-1}^{x+1} c dy dx = c(\frac{3}{2}) + c(8-2) = \frac{15}{2} c = 1$, $c = \frac{2}{15}$

$$E(X) = \frac{2}{15} \int_0^1 \int_0^{x+1} x dy dx + \frac{2}{15} \int_1^4 \int_{x-1}^{x+1} x dy dx = 2.111$$

$$E(Y) = \frac{2}{15} \int_0^1 \int_0^{x+1} y dy dx + \frac{2}{15} \int_1^4 \int_{x-1}^{x+1} y dy dx = 2.156$$

$$\text{Now, } E(XY) = \frac{2}{15} \int_0^1 \int_0^{x+1} xy dy dx + \frac{2}{15} \int_1^4 \int_{x-1}^{x+1} xy dy dx = 5.694$$

$$\sigma_{xy} = 5.694 - (2.111)(2.156) = 1.143$$

$$E(X^2) = 5.678 \quad E(Y^2) = 6.033$$

$$V(x) = 1.222, \quad V(Y) = 1.385$$

$$\rho = \frac{1.143}{\sqrt{1.222} \sqrt{1.385}} = 0.879$$

- 5-37. a) $E(X) = 0$ $E(Y) = 0$

$$E(XY) = \int_1^\infty \int_1^\infty xye^{-x-y} dx dy$$

$$= \int_1^\infty xe^{-x} dx \int_1^\infty ye^{-y} dy$$

$$= E(X)E(Y)$$

Therefore, $\sigma_{XY} = \rho_{XY} = 0$.

- 5-38. $E(X) = 333.33$, $E(Y) = 833.33$

$$E(X^2) = 222,222.2$$

$$V(X) = 222222.2 - (333.33)^2 = 111,113.31$$

$$E(Y^2) = 1,055,556$$

$$V(Y) = 361,117.11$$

$$E(XY) = 6 \times 10^{-6} \int_0^\infty \int_0^\infty xy e^{-0.01x-0.002y} dy dx = 388,888.9$$

$$\sigma_{xy} = 388,888.9 - (333.33)(833.33) = 111,115.01$$

$$\rho = \frac{111,115.01}{\sqrt{111,113.31} \sqrt{361,117.11}} = 0.5547$$

5-39. $E(X) = -1(1/6) + 1(1/6) = 0$

$$E(Y) = -1(1/3) + 1(1/3) = 0$$

$$E(XY) = [-1 \times 0 \times (1/6)] + [-1 \times 0 \times (1/3)] + [1 \times 0 \times (1/3)] + [0 \times 1 \times (1/6)] = 0$$

$$V(X) = 1/3$$

$$V(Y) = 2/3$$

$$\sigma_{xy} = 0 - (0)(0) = 0$$

$$\rho_{xy} = \frac{\sigma_{xy}}{\sigma_x \sigma_y} = \frac{0}{\sqrt{1/3} \sqrt{2/3}} = 0$$

The correlation is zero, but X and Y are not independent, since, for example, if $y = 0$, X must be -1 or 1 .

- 5-40. If X and Y are independent, then $f_{XY}(x, y) = f_X(x)f_Y(y)$ and the range of (X, Y) is rectangular. Therefore,

$$E(XY) = \iint xy f_X(x)f_Y(y) dxdy = \int xf_X(x)dx \int yf_Y(y)dy = E(X)E(Y)$$

hence $\sigma_{xy} = 0$

- 5-41. Suppose the correlation between X and Y is ρ . For constants a , b , c , and d , what is the correlation between the random variables $U = aX+b$ and $V = cY+d$?

Now, $E(U) = aE(X) + b$ and $E(V) = cE(Y) + d$.

Also, $U - E(U) = a[X - E(X)]$ and $V - E(V) = c[Y - E(Y)]$. Then,

$$\sigma_{uv} = E\{[U - E(U)][V - E(V)]\} = acE\{[X - E(X)][Y - E(Y)]\} = ac\sigma_{xy}$$

Also, $\sigma_u^2 = E[U - E(U)]^2 = a^2 E[X - E(X)]^2 = a^2 \sigma_x^2$ and $\sigma_v^2 = c^2 \sigma_y^2$. Then,

$$\rho_{uv} = \frac{ac\sigma_{xy}}{\sqrt{a^2 \sigma_x^2} \sqrt{c^2 \sigma_y^2}} = \begin{cases} \rho_{xy} & \text{if } a \text{ and } c \text{ are of the same sign} \\ -\rho_{xy} & \text{if } a \text{ and } c \text{ differ in sign} \end{cases}$$

Section 5-3

- 5-42. a) board failures caused by assembly defects = $p_1 = 0.5$
 board failures caused by electrical components = $p_2 = 0.4$
 board failures caused by mechanical defects = $p_3 = 0.1$

$$P(X = 5, Y = 3, Z = 2) = \frac{10!}{5!3!2!} 0.5^5 0.4^3 0.1^2 = 0.0504$$

b) Because X is binomial, $P(X = 8) = \binom{10}{8} 0.5^8 0.5^2 = 0.0439$

c) $P(X = 8 | Y = 1) = \frac{P(X = 8, Y = 1)}{P(Y = 1)}$. Now, because $x + y + z = 10$,

$$P(X = 8, Y = 1) = P(X = 8, Y = 1, Z = 1) = \frac{10!}{8!1!1!} 0.5^8 0.4^1 0.1^1 = 0.0141$$

$$P(Y = 1) = \binom{10}{1} 0.4^1 0.6^9 = 0.0403$$

$$P(X = 8 | Y = 1) = \frac{P(X = 8, Y = 1)}{P(Y = 1)} = \frac{0.0141}{0.0403} = 0.3499$$

d) $P(X \geq 8 | Y = 1) = \frac{P(X = 8, Y = 1)}{P(Y = 1)} + \frac{P(X = 9, Y = 1)}{P(Y = 1)}$. Now, because $x + y + z = 10$,

$$P(X = 8, Y = 1) = P(X = 8, Y = 1, Z = 1) = \frac{10!}{8!1!1!} 0.5^8 0.4^1 0.1^1 = 0.0141$$

$$P(X = 9, Y = 1) = P(X = 9, Y = 1, Z = 0) = \frac{10!}{9!1!0!} 0.5^9 0.4^1 0.1^0 = 0.0078$$

$$P(Y = 1) = \binom{10}{1} 0.4^1 0.6^9 = 0.0403$$

$$P(X \geq 8 | Y = 1) = \frac{P(X = 8, Y = 1)}{P(Y = 1)} + \frac{P(X = 9, Y = 1)}{P(Y = 1)} = \frac{0.0141}{0.0403} + \frac{0.0078}{0.0403} = 0.5434$$

e) $P(X = 7, Y = 1 | Z = 2) = \frac{P(X = 7, Y = 1, Z = 2)}{P(Z = 2)}$

$$P(X = 7, Y = 1, Z = 2) = \frac{10!}{7!1!2!} 0.5^7 0.4^1 0.1^2 = 0.0113$$

$$P(Z = 2) = \binom{10}{2} 0.1^2 0.9^8 = 0.1937$$

$$P(X = 7, Y = 1 | Z = 2) = \frac{P(X = 7, Y = 1, Z = 2)}{P(Z = 2)} = \frac{0.0113}{0.1937} = 0.0583$$

- 5-43. a) percentage of slabs classified as high = $p_1 = 0.05$
 percentage of slabs classified as medium = $p_2 = 0.9$
 percentage of slabs classified as low = $p_3 = 0.05$

- b) X is the number of voids independently classified as high $X \geq 0$
 Y is the number of voids independently classified as medium $Y \geq 0$
 Z is the number of voids classified as low and $Z \geq 0$ and $X + Y + Z = 20$

- c) p_1 is the percentage of slabs classified as high.

d) $E(X) = np_1 = 20(0.05) = 1$

$$V(X) = np_1(1 - p_1) = 20(0.05)(0.95) = 0.95$$

- e) $P(X = 1, Y = 17, Z = 3) = 0$ Because the point $(1, 17, 3) \neq 20$ is not in the range of (X, Y, Z) .

f) $P(X \leq 1, Y = 17, Z = 3) = P(X = 0, Y = 17, Z = 3) + P(X = 1, Y = 17, Z = 3)$

$$= \frac{20!}{0!17!3!} 0.05^0 0.9^{17} 0.05^3 + 0 = 0.02376$$

Because the point $(1, 17, 3) \neq 20$ is not in the range of (X, Y, Z) .

g) Because X is binomial, $P(X \leq 1) = \binom{20}{0} 0.05^0 0.95^{20} + \binom{20}{1} 0.05^1 0.95^{19} = 0.7358$

h) Because X is binomial $E(Y) = np = 20(0.9) = 18$

i) The probability is 0 because $x + y + z > 20$

j) $P(X = 2 | Y = 17) = \frac{P(X = 2, Y = 17)}{P(Y = 17)}$. Now, because $x + y + z = 20$,

$$P(X = 2, Y = 17) = P(X = 2, Y = 17, Z = 1) = \frac{20!}{2!17!1!} 0.05^2 0.9^{17} 0.05^1 = 0.0713$$

$$P(X = 2 | Y = 17) = \frac{P(X = 2, Y = 17)}{P(Y = 17)} = \frac{0.0713}{0.1901} = 0.3751$$

k) $E(X | Y = 17) = 0 \left(\frac{P(X = 0, Y = 17)}{P(Y = 17)} \right) + 1 \left(\frac{P(X = 1, Y = 17)}{P(Y = 17)} \right) + 2 \left(\frac{P(X = 2, Y = 17)}{P(Y = 17)} \right) + 3 \left(\frac{P(X = 3, Y = 17)}{P(Y = 17)} \right)$

$$E(X | Y = 17) = 0 \left(\frac{0.02376}{0.1901} \right) + 1 \left(\frac{0.07129}{0.1901} \right) + 2 \left(\frac{0.0713}{0.1901} \right) + 3 \left(\frac{0.02376}{0.1901} \right) = 2$$

- 5-44. a) probability for the first landing page = $p_1 = 0.25$
 probability for the second landing page = $p_2 = 0.25$
 probability for the third landing page = $p_3 = 0.25$
 probability for the fourth landing page = $p_4 = 0.25$

$$P(W = 6, X = 6, Y = 6, Z = 6) = \frac{24!}{6!6!6!} 0.25^6 0.25^6 0.25^6 0.25^6 = 0.0082$$

b) Because $w+x+y+z = 24$ $P(W = 6, X = 6, Y = 6) = P(W = 6, X = 6, Y = 6, Z = 6)$

$$P(W = 6, X = 6, Y = 6) = \frac{24!}{6!6!6!} 0.25^6 0.25^6 0.25^6 0.25^6 = 0.0082$$

c) $P(W = 7, X = 7, Y = 6 | Z = 3) = 0$ Because the point $(7, 7, 6, 3) \neq 24$ is not in the range of (W, X, Y, Z) .

d) $P(W = 8, X = 8, Y = 5 | Z = 3) = \frac{P(W = 8, X = 8, Y = 5, Z = 3)}{P(Z = 3)}$

$$P(W = 8, X = 8, Y = 5, Z = 3) = \frac{24!}{8!8!5!3!} 0.25^8 0.25^8 0.25^5 0.25^3 = 0.0019$$

$$P(Z = 3) = \binom{24}{3} 0.25^3 0.75^{21} = 0.0752$$

$$P(W = 8, X = 8, Y = 5 | Z = 3) = \frac{P(W = 8, X = 8, Y = 5, Z = 3)}{P(Z = 3)} = \frac{0.0019}{0.0752} = 0.0253$$

e) Because W is binomial,

$$P(W \leq 2) = \binom{24}{0} 0.25^0 0.75^{24} + \binom{24}{1} 0.25^1 0.75^{23} + \binom{24}{2} 0.25^2 0.75^{22} = 0.0398$$

f) $E(W) = np = 24(0.25) = 6$

g) $P(W = 6, X = 6) = P(W = 6, X = 6, Y + Z = 12) = \frac{24!}{6!6!12!} 0.25^6 0.25^6 0.5^{12} = 0.0364$

h) $P(W = 6 | X = 6) = \frac{P(W = 6, X = 6)}{P(X = 6)}$

from part g) $P(W = 6, X = 6) = 0.0364$

$$P(X = 6) = \binom{24}{6} 0.25^6 0.75^{18} = 0.1853$$

$$P(W = 6 | X = 6) = \frac{P(W = 6, X = 6)}{P(X = 6)} = \frac{0.0364}{0.1853} = 0.1964$$

- 5-45. a) The probability distribution is multinomial because the result of each trial (a dropped oven) results in either a major, minor or no defect with probability 0.5, 0.4 and 0.1 respectively. Also, the trials are independent
- b) Let X, Y, and Z denote the number of ovens in the sample of four with major, minor, and no defects, respectively.

$$P(X = 2, Y = 2, Z = 0) = \frac{4!}{2!2!0!} 0.5^2 0.4^2 0.1^0 = 0.24$$

$$\text{c) } P(X = 0, Y = 0, Z = 4) = \frac{4!}{0!0!4!} 0.5^0 0.4^0 0.1^4 = 0.0001$$

d) $f_{XY}(x,y) = \sum_R f_{XYZ}(x,y,z)$ where R is the set of values for z such that $x+y+z = 4$. That is, R consists of the single value $z = 4-x-y$ and

$$f_{XY}(x, y) = \frac{4!}{x! y! (4-x-y)!} 0.5^x 0.4^y 0.1^{4-x-y} \quad \text{for } x + y \leq 4.$$

$$\text{e) } E(X) = np_1 = 4(0.5) = 2$$

$$\text{f) } E(Y) = np_2 = 4(0.4) = 1.6$$

$$\text{g) } P(X = 2 | Y = 2) = \frac{P(X = 2, Y = 2)}{P(Y = 2)} = \frac{0.24}{0.3456} = 0.694$$

$$P(Y = 2) = \binom{4}{2} 0.4^2 0.6^2 = 0.3456 \text{ from the binomial marginal distribution of } Y$$

h) Not possible, $x+y+z = 4$, the probability is zero.

$$\text{i) } P(X | Y = 2) = P(X = 0 | Y = 2), P(X = 1 | Y = 2), P(X = 2 | Y = 2)$$

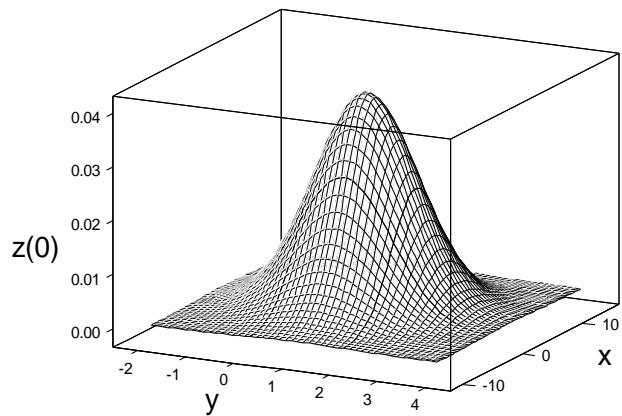
$$P(X = 0 | Y = 2) = \frac{P(X = 0, Y = 2)}{P(Y = 2)} = \left(\frac{4!}{0!2!2!} 0.5^0 0.4^2 0.1^2 \right) / 0.3456 = 0.0278$$

$$P(X = 1 | Y = 2) = \frac{P(X = 1, Y = 2)}{P(Y = 2)} = \left(\frac{4!}{1!2!1!} 0.5^1 0.4^2 0.1^1 \right) / 0.3456 = 0.2778$$

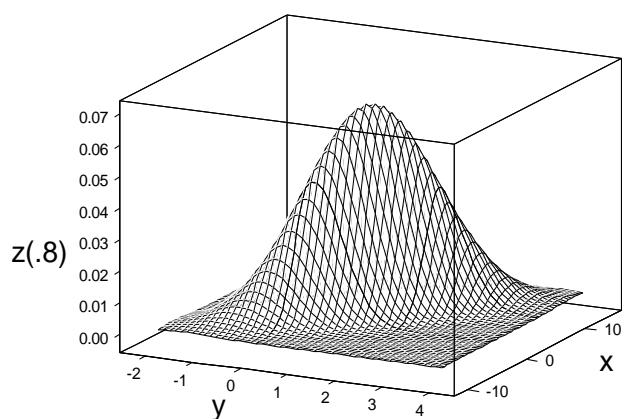
$$P(X = 2 | Y = 2) = \frac{P(X = 2, Y = 2)}{P(Y = 2)} = \left(\frac{4!}{2!2!0!} 0.5^2 0.4^2 0.1^0 \right) / 0.3456 = 0.6944$$

$$\text{j) } E(X|Y=2) = 0(0.0278) + 1(0.2778) + 2(0.6944) = 1.6666$$

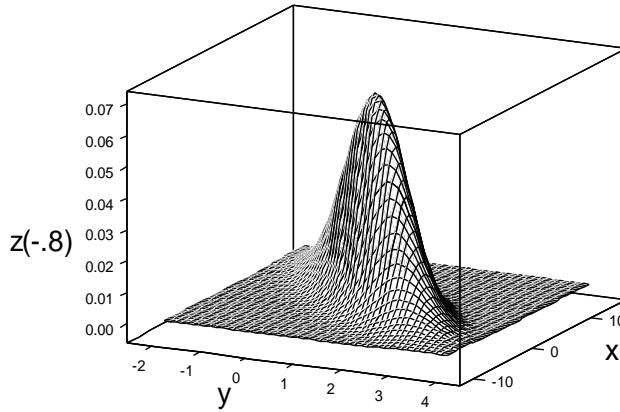
- 5-46. a)



b)



c)



- 5-47. Because $\rho = 0$ and X and Y are normally distributed, X and Y are independent. Therefore,
- (a) $P(2.95 < X < 3.05) = P\left(\frac{2.95-3}{0.04} < Z < \frac{3.05-3}{0.04}\right) = 0.7887$
 - (b) $P(7.60 < Y < 7.80) = P\left(\frac{7.6-8.0}{0.08} < Z < \frac{7.8-8.0}{0.08}\right) = 0.00621$
 - (c) $P(2.95 < X < 3.05, 7.60 < Y < 7.80) = P(2.95 < X < 3.05) P(7.60 < Y < 7.80) = P\left(\frac{2.95-3}{0.04} < Z < \frac{3.05-3}{0.04}\right) P\left(\frac{7.6-8.0}{0.08} < Z < \frac{7.8-8.0}{0.08}\right) = (0.7887)(0.00621) = 0.0049$
- 5-48.
- (a) $\rho = \text{cov}(X, Y)/\sigma_x\sigma_y = 0.6$; $\text{cov}(X, Y) = 0.6 * 3 * 5 = 9$
 - (b) The marginal probability distribution of X is normal with mean μ_x , σ_x .
 - (c) $P(X < 116) = P(X-120 < -4) = P((X-120)/5 < -0.8) = P(Z < -0.8) = 0.21$
 - (d) The conditional probability distribution of X given $Y=102$ is bivariate normal distribution with mean and variance

$$\mu_{X|y=102} = 120 - 100 * 0.6 * (5/3) + (5/3) * 0.6 * 102 = 122$$

$$\sigma_{X|y=102}^2 = 25(1-0.36) = 16$$
 - (e) $P(X < 116|Y=102) = P(Z < (116-122)/4) = 0.0668$
- 5-49.
- Because $\rho = 0$ and X and Y are normally distributed, X and Y are independent. Therefore, $\mu_X = 0.1$ mm, $\sigma_X = 0.00031$ mm, $\mu_Y = 0.23$ mm, $\sigma_Y = 0.00017$ mm
 - Probability X is within specification limits is
- $$P(0.099535 < X < 0.100465) = P\left(\frac{0.099535 - 0.1}{0.00031} < Z < \frac{0.100465 - 0.1}{0.00031}\right)$$
- $$= P(-1.5 < Z < 1.5) = P(Z < 1.5) - P(Z < -1.5) = 0.8664$$
- Probability that Y is within specification limits is
- $$P(0.22966 < X < 0.23034) = P\left(\frac{0.22966 - 0.23}{0.00017} < Z < \frac{0.23034 - 0.23}{0.00017}\right)$$
- $$= P(-2 < Z < 2) = P(Z < 2) - P(Z < -2) = 0.9545$$
- Probability that a randomly selected lamp is within specification limits is $(0.8664)(0.9545) = 0.8270$

- 5-50. a) By completing the square in the numerator of the exponent of the bivariate normal PDF, the joint PDF can be written as

$$f_{Y|X=x} = \frac{f_{XY}(x, y)}{f_X(x)} = \frac{\frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}}e^{-\frac{\left[\frac{1}{\sigma_y^2}\left(y-(\mu_Y+\rho\frac{\sigma_y}{\sigma_x}(x-\mu_x))\right)^2+(1-\rho^2)\left(\frac{x-\mu_x}{\sigma_x}\right)^2\right]}{2(1-\rho^2)}}}{\frac{1}{\sqrt{2\pi\sigma_x}}e^{-\frac{\left[\frac{x-\mu_x}{\sigma_x}\right]^2}{2}}}$$

Also, $f_X(x) = \frac{1}{\sqrt{2\pi\sigma_x}}e^{-\frac{\left[\frac{x-\mu_x}{\sigma_x}\right]^2}{2}}$ By definition,

$$f_{Y|X=x} = \frac{f_{XY}(x, y)}{f_X(x)} = \frac{\frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}}e^{-\frac{\left[\frac{1}{\sigma_y^2}\left(y-(\mu_Y+\rho\frac{\sigma_y}{\sigma_x}(x-\mu_x))\right)^2+(1-\rho^2)\left(\frac{x-\mu_x}{\sigma_x}\right)^2\right]}{2(1-\rho^2)}}}{\frac{1}{\sqrt{2\pi\sigma_x}}e^{-\frac{\left[\frac{x-\mu_x}{\sigma_x}\right]^2}{2(1-\rho^2)}}}$$

$$= \frac{1}{\sqrt{2\pi\sigma_y\sqrt{1-\rho^2}}}e^{-\frac{\left[\frac{1}{\sigma_y^2}\left(y-(\mu_Y+\rho\frac{\sigma_y}{\sigma_x}(x-\mu_x))\right)^2\right]}{2(1-\rho^2)}}$$

Now $f_{Y|X=x}$ is in the form of a normal distribution.

b) $E(Y|X=x) = \mu_y + \rho\frac{\sigma_y}{\sigma_x}(x - \mu_x)$. This answer can be seen from part a). Because the PDF is in the form of a normal distribution, then the mean can be obtained from the exponent.

c) $V(Y|X=x) = \sigma_y^2(1-\rho^2)$. This answer can be seen from part a). Because the PDF is in the form of a normal distribution, then the variance can be obtained from the exponent.

- 5-51.

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XY}(x, y) dx dy = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \left[\frac{1}{2\pi\sigma_x\sigma_y} e^{-\frac{1}{2}\left[\frac{(x-\mu_x)^2}{\sigma_x^2} + \frac{(y-\mu_y)^2}{\sigma_y^2}\right]} \right] dx dy =$$

$$\int_{-\infty}^{\infty} \left[\frac{1}{\sqrt{2\pi\sigma_x}} e^{-\frac{1}{2}\left[\frac{(x-\mu_x)^2}{\sigma_x^2}\right]} \right] dx \int_{-\infty}^{\infty} \left[\frac{1}{\sqrt{2\pi\sigma_y}} e^{-\frac{1}{2}\left[\frac{(y-\mu_y)^2}{\sigma_y^2}\right]} \right] dy$$

and each of the last two integrals is recognized as the integral of a normal probability density function from $-\infty$ to ∞ . That is, each integral equals one. Because $f_{XY}(x, y) = f(x)f(y)$, X and Y are independent.

- 5-52.

$$\text{Let } f_{XY}(x, y) = \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\left[\left(\frac{X-\mu_X}{\sigma_X}\right)^2 - \frac{2\rho(X-\mu_X)(Y-\mu_Y)}{\sigma_X\sigma_Y} + \left(\frac{Y-\mu_Y}{\sigma_Y}\right)^2\right] / 2(1-\rho^2)}$$

Completing the square in the numerator of the exponent we get:

$$\left[\left(\frac{X-\mu_X}{\sigma_X}\right)^2 - \frac{2\rho(X-\mu_X)(Y-\mu_Y)}{\sigma_X\sigma_Y} + \left(\frac{Y-\mu_Y}{\sigma_Y}\right)^2\right] = \left[\left(\frac{Y-\mu_Y}{\sigma_Y}\right) - \rho\left(\frac{X-\mu_X}{\sigma_X}\right)\right]^2 + (1-\rho^2)\left(\frac{X-\mu_X}{\sigma_X}\right)^2$$

But,

$$\left(\frac{Y-\mu_Y}{\sigma_Y}\right) - \rho\left(\frac{X-\mu_X}{\sigma_X}\right) = \frac{1}{\sigma_Y} \left[(Y-\mu_Y) - \rho \frac{\sigma_Y}{\sigma_X} (X-\mu_X) \right] = \frac{1}{\sigma_Y} \left[(Y - (\mu_Y + \rho \frac{\sigma_Y}{\sigma_X} (X-\mu_X))) \right]$$

Substituting into $f_{XY}(x, y)$, we get

$$\begin{aligned} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XY}(x, y) &= \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\left[\frac{1}{\sigma_Y^2} \left[y - (\mu_Y + \rho \frac{\sigma_Y}{\sigma_X} (x-\mu_X)) \right]^2 + (1-\rho^2) \left(\frac{x-\mu_X}{\sigma_X} \right)^2 \right] / 2(1-\rho^2)} dy dx \\ &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma_x} e^{-\frac{1}{2} \left(\frac{x-\mu_X}{\sigma_X} \right)^2} dx \times \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma_y\sqrt{1-\rho^2}} e^{-\left[\frac{\left(y - (\mu_Y + \rho \frac{\sigma_Y}{\sigma_X} (x-\mu_X)) \right)^2}{2\sigma_Y^2(1-\rho^2)} \right]} dy \end{aligned}$$

The integrand in the second integral above is in the form of a normally distributed random variable. By definition of the integral over this function, the second integral is equal to 1:

$$\begin{aligned} &\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma_x} e^{-\frac{1}{2} \left(\frac{x-\mu_X}{\sigma_X} \right)^2} dx \times \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma_y\sqrt{1-\rho^2}} e^{-\left[\frac{\left(y - (\mu_Y + \rho \frac{\sigma_Y}{\sigma_X} (x-\mu_X)) \right)^2}{2\sigma_Y^2(1-\rho^2)} \right]} dy \\ &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma_x} e^{-\frac{1}{2} \left(\frac{x-\mu_X}{\sigma_X} \right)^2} dx \times 1 \end{aligned}$$

The remaining integral is also the integral of a normally distributed random variable and therefore, it also integrates to 1, by definition. Therefore,

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XY}(x, y) = 1$$

$$\begin{aligned}
 f_X(x) &= \int_{-\infty}^{\infty} \left[\frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{0.5}{1-\rho^2} \left[\frac{(x-\mu_x)^2}{\sigma_x^2} - \frac{2\rho(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y} + \frac{(y-\mu_y)^2}{\sigma_y^2} \right]} \right] dy \\
 &= \frac{1}{\sqrt{2\pi}\sigma_x} e^{-\frac{0.5(x-\mu_x)^2}{\sigma_x^2}} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma_y\sqrt{1-\rho^2}} e^{-\frac{0.5}{1-\rho^2} \left[\left(\frac{(y-\mu_y)}{\sigma_y} - \frac{\rho(x-\mu_x)}{\sigma_x} \right)^2 - \left(\frac{\rho(x-\mu_x)}{\sigma_x} \right)^2 \right]} dy \\
 &= \frac{1}{\sqrt{2\pi}\sigma_x} e^{-\frac{0.5(x-\mu_x)^2}{\sigma_x^2}} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma_y\sqrt{1-\rho^2}} e^{-\frac{0.5}{1-\rho^2} \left(\frac{(y-\mu_y)}{\sigma_y} - \frac{\rho(x-\mu_x)}{\sigma_x} \right)^2} dy
 \end{aligned}$$

The last integral is recognized as the integral of a normal probability density with mean $\mu_y + \frac{\sigma_y \rho(x-\mu_x)}{\sigma_x}$ and variance $\sigma_y^2(1-\rho^2)$. Therefore, the last integral equals one and the requested result is obtained.

Section 5-4

5-54. a) $E(2X + 3Y) = 2(0) + 3(10) = 30$

b) $V(2X + 3Y) = 4V(X) + 9V(Y) = 101$

c) $2X + 3Y$ is normally distributed with mean 30 and variance 101. Therefore,

$$P(2X + 3Y < 30) = P(Z < \frac{30-30}{\sqrt{101}}) = P(Z < 0) = 0.5$$

d) $P(2X + 3Y < 40) = P(Z < \frac{40-30}{\sqrt{101}}) = P(Z < 0.995) = 0.8389$

5-55. a) $E(3X+2Y) = 3*2+2*6=18$

b) $V(3X+2Y) = 9*5+4*9 = 81$

c) $3X+2Y \sim N(18, 81)$

$$P(3X+2Y < 18) = P(Z < (18-18)/81^{0.5}) = 0.5$$

d) $P(3X+2Y < 28) = P(Z < (28-18)/81^{0.5}) = P(Z < 1.1111) = 0.8665$

5-56. $Y = 10X$ and $E(Y) = 10E(X) = 90\text{mm}$.

$$V(Y) = 10^2 V(X) = 40\text{mm}^2$$

5-57. a) Let T denote the total thickness. Then, $T = X + Y$ and $E(T) = 6\text{ mm}$,

$$V(T) = 0.1^2 + 0.1^2 = 0.02\text{mm}^2, \text{ and } \sigma_T = 0.1414\text{ mm.}$$

b)

$$P(T > 6.3) = P\left(Z > \frac{6.3-6}{0.1414}\right) = P(Z > 2.12)$$

$$= 1 - P(Z < 2.12) = 1 - 0.983 = 0.0170$$

5-58. a) X : time of wheel throwing. $X \sim N(40, 4)$

Y : time of wheel firing. $Y \sim N(60, 9)$

$X + Y \sim N(100, 13)$

$$P(X + Y \leq 90) = P(Z < (90 - 100)/13^{0.5}) = P(Z < -2.774) = 0.0028$$

(b) $P(X + Y > 110) = 1 - P(Z < (110 - 100)/13^{0.5}) = 1 - P(Z < 2.774) = 1 - 0.9972 = 0.0028$

5-59. a) $X \sim N(0.1, 0.00031)$ and $Y \sim N(0.23, 0.00017)$ Let T denote the total thickness.

Then, $T = X + Y$ and $E(T) = 0.33$ mm,

$$V(T) = 0.00031^2 + 0.00017^2 = 1.25 \times 10^{-7} \text{ mm}^2, \text{ and } \sigma_T = 0.000354 \text{ mm.}$$

$$P(T < 0.2340) = P\left(Z < \frac{0.2340 - 0.33}{0.000354}\right) = P(Z < -271.2) \approx 0$$

$$\text{b)} \quad P(T > 0.2405) = P\left(Z > \frac{0.2405 - 0.33}{0.000354}\right) = P(Z > -253) = 1 - P(Z < 253) \approx 1$$

- 5-60. Let D denote the width of the casing minus the width of the door. Then, D is normally distributed.

$$\text{a)} \quad E(D) = 0.4 \quad V(D) = (0.4)^2 + (0.2)^2 = 0.2 \quad \sigma_D = \sqrt{0.2} = 0.4472$$

$$\text{b)} \quad P(D > 0.8) = P\left(Z > \frac{0.8 - 0.4}{\sqrt{0.2}}\right) = P(Z > 0.89) = 0.187$$

$$\text{c)} \quad P(D < 0) = P\left(Z < \frac{0 - 0.4}{\sqrt{0.2}}\right) = P(Z < -0.89) = 0.187$$

- 5-61. X = time of ACL reconstruction surgery for high-volume hospitals.

$$X \sim N(129, 196)$$

$$E(X_1 + X_2 + \dots + X_{10}) = 8 * 129 = 1032$$

$$V(X_1 + X_2 + \dots + X_{10}) = 64 * 196 = 8256$$

- 5-62. a) Let \bar{X} denote the average fill-volume of 100 cans. $\sigma_{\bar{X}} = \sqrt{\frac{15^2}{100}} = 1.5$.

$$\text{b)} \quad E(\bar{X}) = 358 \text{ and } P(\bar{X} < 355) = P\left(Z < \frac{355 - 358}{1.5}\right) = P(Z < -2) = 0.023$$

$$\text{c)} \quad P(\bar{X} < 355) = 0.005 \text{ implies that } P\left(Z < \frac{355 - \mu}{1.5}\right) = 0.005.$$

$$\text{Then } \frac{355 - \mu}{1.5} = -2.58 \text{ and } \mu = 358.87.$$

$$\text{d)} \quad P(\bar{X} < 355) = 0.005 \text{ implies that } P\left(Z < \frac{355 - 358}{\sigma/\sqrt{100}}\right) = 0.005.$$

$$\text{Then } \frac{355 - 358}{\sigma/\sqrt{100}} = -2.58 \text{ and } \sigma = 11.628.$$

$$\text{e)} \quad P(\bar{X} < 355) = 0.01 \text{ implies that } P\left(Z < \frac{355 - 358}{15/\sqrt{n}}\right) = 0.01.$$

$$\text{Then } \frac{355 - 358}{15/\sqrt{n}} = -2.33 \text{ and } n = 135.72 \approx 136.$$

- 5-63. Let \bar{X} denote the average thickness of 10 wafers. Then, $E(\bar{X}) = 10$ and $V(\bar{X}) = 0.1$.

$$\text{a)} \quad P(8 < \bar{X} < 12) = P\left(\frac{8-10}{\sqrt{0.1}} < Z < \frac{12-10}{\sqrt{0.1}}\right) = P(-6.32 < Z < 6.32) \approx 1.$$

The answer is ≈ 0

$$\text{b)} \quad P(\bar{X} > 11) = 0.05 \text{ and } \sigma_{\bar{X}} = \sqrt{\frac{1}{n}}.$$

Therefore, $P(\bar{X} > 11) = P(Z > \frac{11-10}{\sqrt{\frac{1}{10}}}) = 0.05$, $\frac{11-10}{\sqrt{\frac{1}{10}}} = 1.65$ and $n = 2.72$ which is rounded up to 3.

$$\text{c)} \quad P(\bar{X} > 11) = 0.0005 \text{ and } \sigma_{\bar{X}} = \sqrt{\frac{1}{10}}.$$

$$\text{Therefore, } P(\bar{X} > 11) = P(Z > \frac{11-10}{\sqrt{\frac{1}{10}}}) = 0.0005, \frac{11-10}{\sqrt{\frac{1}{10}}} = 3.29$$

$$\sigma = \sqrt{10} / 3.29 = 0.9612$$

5-64. $X \sim N(75, 225)$

a) Let $Y = X_1 + X_2 + \dots + X_{25}$, $E(Y) = 25E(X) = 1875$, $V(Y) = 25^2(225) = 140625$

$$P(Y > 1950) = P\left(Z > \frac{1950 - 1875}{\sqrt{140625}}\right) = P(Z > 0.2) = 1 - P(Z < 0.2) = 1 - 0.5793 = 0.4207$$

$$b) P(Y > x) = 0.0002 \text{ implies that } P\left(Z > \frac{x - 1875}{\sqrt{140625}}\right) = 0.0002.$$

$$\text{Then } \frac{x - 1875}{375} = 3.54 \text{ and } x = 3202.5$$

5-65. W: weights of parts E: measurement error.

$W \sim N(\mu_w, \sigma_w^2)$, $E \sim N(0, \sigma_e^2)$, $W+E \sim N(\mu_w, \sigma_w^2 + \sigma_e^2)$.

W_{sp} = weights of the specification P

$$(a) P(W > \mu_w + 3\sigma_w) + P(W < \mu_w - 3\sigma_w) = P(Z > 3) + P(Z < -3) = 0.0027$$

$$(b) P(W+E > \mu_w + 3\sigma_w) + P(W+E < \mu_w - 3\sigma_w) \\ = P(Z > 3\sigma_w / (\sigma_w^2 + \sigma_e^2)^{1/2}) + P(Z < -3\sigma_w / (\sigma_w^2 + \sigma_e^2)^{1/2})$$

Because $\sigma_e^2 = 0.25\sigma_w^2$ the probability is

$$= P(Z > 3\sigma_w / (1.25\sigma_w^2)^{1/2}) + P(Z < -3\sigma_w / (1.25\sigma_w^2)^{1/2}) \\ = P(Z > 2.68) + P(Z < -2.68) = 2(0.003681) = 0.0074$$

No.

$$(c) P(E + \mu_w + 2\sigma_w > \mu_w + 3\sigma_w) = P(E > \sigma_w) = P(Z > \sigma_w / (0.25\sigma_w^2)^{1/2}) = P(Z > 2) = 0.0228$$

$$\text{Also, } P(E + \mu_w + 2\sigma_w < \mu_w - 3\sigma_w) = P(E < -5\sigma_w) = P(Z < -5\sigma_w / (0.25\sigma_w^2)^{1/2}) = P(Z < -10) \approx 0$$

5-66. $D = A - B - C$

$$a) E(D) = 10 - 2 - 2 = 6 \text{ mm}$$

$$V(D) = 0.1^2 + 0.1^2 + 0.1^2 = 0.03 \text{ mm}^2$$

$$\sigma_D = 0.1732 \text{ mm}$$

$$b) P(D < 5.9) = P\left(Z < \frac{5.9 - 6}{0.1732}\right) = P(Z < -0.58) = 0.281.$$

Section 5-5

$$5-67. f_Y(y) = \frac{1}{4} \text{ at } y = 5, 8, 11, 14$$

5-68. Because $X \geq 0$, the transformation is one-to-one; that is $y = x^3$ and $x = \sqrt[3]{y}$. From equation 5-30,

$$f_Y(y) = f_X(\sqrt[3]{y}) = \binom{3}{\sqrt[3]{y}} p^{\sqrt[3]{y}} (1-p)^{3-\sqrt[3]{y}} \text{ for } y = 0, 1, 8, 27.$$

$$\text{If } p = 0.25, f_Y(y) = \binom{3}{\sqrt[3]{y}} (0.25)^{\sqrt[3]{y}} (0.75)^{3-\sqrt[3]{y}} \text{ for } y = 0, 1, 8, 27.$$

$$5-69. a) f_Y(y) = f_X\left(\frac{y-10}{2}\right)\left(\frac{1}{2}\right) = \frac{y-10}{96} \text{ for } 10 \leq y \leq 22$$

$$b) E(Y) = \int_{10}^{22} \frac{y^2 - 10y}{96} dy = \frac{1}{96} \left[\frac{y^3}{3} - \frac{10y^2}{2} \right]_{10}^{22} = 13.5$$

- 5-70. Because $y = -2 \ln x$, $e^{-\frac{y}{2}} = x$. Then, $f_Y(y) = f_X(e^{-\frac{y}{2}}) \left| -\frac{1}{2} e^{-\frac{y}{2}} \right| = \frac{1}{2} e^{-\frac{y}{2}}$ for $0 \leq e^{-\frac{y}{2}} \leq 1$ or $y \geq 0$, which is an exponential distribution with $\lambda = 1/2$ (which equals a chi-square distribution with $k = 2$ degrees of freedom).

- 5-71. a) If $y = x^2$, then $x = \sqrt{y}$ for $x \geq 0$ and $y \geq 0$. Thus, $f_Y(y) = f_X(\sqrt{y}) \frac{1}{2} y^{-\frac{1}{2}} = \frac{e^{-\sqrt{y}}}{2\sqrt{y}}$ for $y > 0$.
 b) If $y = x^{1/3}$, then $x = y^3$ for $x \geq 0$ and $y \geq 0$. Thus, $f_Y(y) = f_X(y^3) 3y = 3ye^{-y^3}$ for $y > 0$.
 c) If $y = \ln x$, then $x = e^y$ for $x \geq 0$. Thus, $f_Y(y) = f_X(e^y) e^y = e^y e^{-e^y} = e^{y-e^y}$ for $-\infty < y < \infty$.

- 5-72. a) Now, $\int_0^\infty av^2 e^{-bv} dv$ must equal one. Let $u = bv$, then $1 = a \int_0^\infty (\frac{u}{b})^2 e^{-u} \frac{du}{b} = \frac{a}{b^3} \int_0^\infty u^2 e^{-u} du$. From the definition of the gamma function the last expression is $\frac{a}{b^3} \Gamma(3) = \frac{2a}{b^3}$. Therefore,
- $$a = \frac{b^3}{2}.$$

b) If $w = \frac{mv^2}{2}$, then $v = \sqrt{\frac{2w}{m}}$ for $v \geq 0$, $w \geq 0$.

$$\begin{aligned} f_w(w) &= f_v\left(\sqrt{\frac{2w}{m}}\right) \frac{dv}{dw} = \frac{b^3 2w}{2m} e^{-b\sqrt{\frac{2w}{m}}} (2mw)^{-1/2} \\ &= \frac{b^3 m^{-3/2}}{\sqrt{2}} w^{1/2} e^{-b\sqrt{\frac{2w}{m}}} \end{aligned}$$

for $w \geq 0$.

- 5-73. If $y = e^{x+1}$, then $x = \ln y - 1$ for $1 \leq x \leq 2$ and $e^2 \leq y \leq e^3$. Thus,

$$f_Y(y) = f_X(\ln y - 1) \frac{1}{y} = \frac{1}{y} \text{ for } 2 \leq \ln y \leq 3. \text{ That is, } f_Y(y) = \frac{1}{y} \text{ for } e^2 \leq y \leq e^3.$$

- 5-74. If $y = (x-2)^2$, then $x = 2 - \sqrt{y}$ for $0 \leq x \leq 2$ and $x = 2 + \sqrt{y}$ for $2 \leq x \leq 4$. Thus,

$$\begin{aligned} f_Y(y) &= f_X(2 - \sqrt{y}) \left| -\frac{1}{2} y^{-1/2} \right| + f_X(2 + \sqrt{y}) \left| \frac{1}{2} y^{-1/2} \right| \\ &= \frac{2 - \sqrt{y}}{24\sqrt{y}} + \frac{2 + \sqrt{y}}{24\sqrt{y}} \\ &= \left(\frac{1}{6}\right) y^{-1/2} \text{ for } 0 \leq y \leq 4 \end{aligned}$$

Section 5-6

- 5-75.

$$\text{a) } M_X(t) = E(e^{tX}) = \sum_{x=1}^m \frac{1}{m} e^{tx} = \frac{1}{m} \sum_{x=0}^{m-1} e^{t(x+1)} = \frac{e^t}{m} \sum_{x=0}^{m-1} e^{tx} = \frac{e^t(1-e^{tm})}{m(1-e^t)}$$

$$\text{b) } E(X) = \mu = \mu_1' = \left. \frac{dM_X(t)}{dt} \right|_{t=0} = \frac{m+1}{2}$$

$$V(X) = \sigma^2 = E(X^2) - [E(X)]^2 = \mu_2' - \mu^2 = \frac{(m+1)(2m+1)}{6} - \left(\frac{m+1}{2} \right)^2 = \frac{m^2 - 1}{12}$$

5-76.

$$\text{a) } M_X(t) = E(e^{tX}) = \sum_{x=0}^{\infty} e^{tx} \frac{e^{-\lambda} \lambda^x}{x!} = e^{-\lambda} \sum_{x=0}^{\infty} \frac{(\lambda e^t)^x}{x!} = e^{-\lambda} (e^{\lambda e^t}) = e^{\lambda(e^t - 1)}$$

$$\text{b) } E(X) = \mu = \mu_1' = \left. \frac{dM_X(t)}{dt} \right|_{t=0} = \lambda e^t (e^{\lambda(e^t - 1)}) \Big|_{t=0} = \lambda$$

$$V(X) = \sigma^2 = E(X^2) - [E(X)]^2 = \mu_2' - \mu^2$$

$$\mu_2' = \left. \frac{d^2M_X(t)}{dt^2} \right|_{t=0} = \left[\lambda e^t (e^{\lambda(e^t - 1)}) + (\lambda e^t)^2 (e^{\lambda(e^t - 1)}) \right] \Big|_{t=0} = \lambda + \lambda^2$$

$$\sigma^2 = \mu_2' - \mu^2 = \lambda + \lambda^2 - \lambda^2 = \lambda$$

$$\text{5-77. } M_X(t) = E(e^{tX}) = \sum_{x=1}^{\infty} e^{tx} p(1-p)^{x-1} = \frac{p}{1-p} \sum_{x=1}^{\infty} (e^t(1-p))^x$$

$$= \frac{p}{1-p} \left(\frac{e^t(1-p)}{1-e^t(1-p)} \right) = \frac{pe^t}{1-(1-p)e^t}$$

$$E(X) = \mu = \mu_1' = \left. \frac{dM_X(t)}{dt} \right|_{t=0} = \left. \frac{pe^t(1-(1-p)e^t) - pe^t(-(1-p)e^t)}{(1-(1-p)e^t)^2} \right|_{t=0}$$

$$= \left. \frac{pe^t}{(1-(1-p)e^t)^2} \right|_{t=0} = \frac{p}{p^2} = \frac{1}{p}$$

$$V(X) = \sigma^2 = E(X^2) - [E(X)]^2 = \mu_2' - \mu^2$$

$$\mu_2' = \left. \frac{d^2M_X(t)}{dt^2} \right|_{t=0} = \left. \frac{pe^t(1-(1-p)e^t)^2 - pe^t(2)(1-(1-p)e^t)(-(1-p)e^t)}{(1-(1-p)e^t)^4} \right|_{t=0}$$

$$= \left. \frac{pe^t(1+(1-p)e^t)}{(1-(1-p)e^t)^3} \right|_{t=0} = \frac{2-p}{p^2}$$

$$\sigma^2 = \mu_2' - \mu^2 = \frac{2-p}{p^2} - \frac{1}{p^2} = \frac{1-p}{p^2}$$

$$\text{5-78. } M_Y(t) = M_{X_1}(t)M_{X_2}(t) = ((1-2t)^{-k_1/2})(1-2t)^{-k_2/2}) = (1-2t)^{-(k_1+k_2)/2}$$

As a result, Y is a chi-squared random variable with $k_1 + k_2$ degrees of freedom.

5-79

$$a) M_X(t) = E(e^{tX}) = \int_{x=0}^{\infty} e^{tx} (4xe^{-2x}) dx = 4 \int_{x=0}^{\infty} xe^{(t-2)x} dx$$

Using integration by parts, we have:

$$M_X(t) = 4 \lim_{c \rightarrow \infty} \left[\frac{xe^{(t-2)x}}{t-2} - \frac{e^{(t-2)x}}{(t-2)^2} \right]_0^c = 4 \lim_{c \rightarrow \infty} \left[\left(\frac{x}{t-2} - \frac{1}{(t-2)^2} \right) e^{(t-2)x} \right]_0^c = \frac{4}{(t-2)^2}$$

$$b) E(X) = \mu = \mu'_1 = \frac{dM_X(t)}{dt} \Big|_{t=0} = \frac{-8}{(t-2)^3} \Big|_{t=0} = 1$$

$$\mu'_2 = \frac{d^2M_X(t)}{dt^2} \Big|_{t=0} = \frac{24}{(t-2)^4} \Big|_{t=0} = 1.5$$

$$V(X) = \sigma^2 = E(X^2) - [E(X)]^2 = \mu'_2 - \mu^2 = 1.5 - 1 = 0.5$$

5-80.

$$a) M_X(t) = E(e^{tX}) = \int_{\alpha}^{\beta} e^{tx} \frac{1}{\beta-\alpha} dx = \frac{1}{\beta-\alpha} \left(\frac{e^{tx}}{t} \right) \Big|_{x=\alpha}^{x=\beta} = \frac{1}{\beta-\alpha} \left(\frac{e^{t\beta}}{t} - \frac{e^{t\alpha}}{t} \right) = \frac{e^{t\beta} - e^{t\alpha}}{t(\beta-\alpha)}$$

$$b) E(X) = \mu = \mu'_1 = \frac{dM_X(t)}{dt} \Big|_{t=0}$$

$$\frac{dM_X(t)}{dt} = \frac{(\beta e^{t\beta} - \alpha e^{t\alpha})(t(\beta-\alpha)) - (e^{t\beta} - e^{t\alpha})(\beta-\alpha)}{t^2(\beta-\alpha)^2} = \frac{t(\beta e^{t\beta} - \alpha e^{t\alpha}) - e^{t\beta} + e^{t\alpha}}{t^2(\beta-\alpha)}$$

$\frac{dM_X(t)}{dt}$ is undefined at $t = 0$ since there is t^2 in the denominator. Indeed, it has an

indeterminate form of $\frac{0}{0}$ when it is evaluated at $t = 0$. As a result, we need to use L'Hopital's rule and differentiate the numerator and denominator.

$$E(X) = \lim_{t \rightarrow 0} \frac{dM_X(t)}{dt} = \lim_{t \rightarrow 0} \frac{t(\beta e^{t\beta} - \alpha e^{t\alpha}) - e^{t\beta} + e^{t\alpha}}{t^2(\beta-\alpha)}$$

$$= \lim_{t \rightarrow 0} \frac{(\beta e^{t\beta} - \alpha e^{t\alpha}) + t(\beta^2 e^{t\beta} - \alpha^2 e^{t\alpha}) - \beta e^{t\beta} + \alpha e^{t\alpha}}{2t(\beta-\alpha)}$$

$$= \lim_{t \rightarrow 0} \frac{\beta^2 e^{t\beta} - \alpha^2 e^{t\alpha}}{2(\beta-\alpha)} = \frac{\beta^2 - \alpha^2}{2(\beta-\alpha)} = \frac{\alpha + \beta}{2}$$

$$\mu'_2 = \frac{d^2M_X(t)}{dt^2} \Big|_{t=0}$$

$$\begin{aligned}\frac{d^2M_X(t)}{dt^2} &= \frac{t^3(\beta^2 e^{t\beta} - \alpha^2 e^{t\alpha})(\beta - \alpha) - 2t(\beta - \alpha)(t(\beta e^{t\beta} - \alpha e^{t\alpha}) - e^{t\beta} + e^{t\alpha})}{t^4(\beta - \alpha)^2} \\ &= \frac{t^2(\beta^2 e^{t\beta} - \alpha^2 e^{t\alpha}) - 2t(\beta e^{t\beta} - \alpha e^{t\alpha}) + 2e^{t\beta} - 2e^{t\alpha}}{t^3(\beta - \alpha)}\end{aligned}$$

$\frac{d^2M_X(t)}{dt^2}$ has the same indefinite form of $\frac{0}{0}$ when it is evaluated at $t = 0$. We need to use L'Hopital's rule again.

$$\begin{aligned}\mu_2' &= \lim_{t \rightarrow 0} \frac{t^2(\beta^2 e^{t\beta} - \alpha^2 e^{t\alpha}) - 2t(\beta e^{t\beta} - \alpha e^{t\alpha}) + 2e^{t\beta} - 2e^{t\alpha}}{t^3(\beta - \alpha)} \\ &= \lim_{t \rightarrow 0} \frac{2t(\beta^2 e^{t\beta} - \alpha^2 e^{t\alpha}) + t^2(\beta^3 e^{t\beta} - \alpha^3 e^{t\alpha}) - 2(\beta e^{t\beta} - \alpha e^{t\alpha}) - 2t(\beta^2 e^{t\beta} - \alpha^2 e^{t\alpha}) + 2\beta e^{t\beta} - 2\alpha e^{t\alpha}}{3t^2(\beta - \alpha)} \\ &= \lim_{t \rightarrow 0} \frac{t^2(\beta^3 e^{t\beta} - \alpha^3 e^{t\alpha})}{3t^2(\beta - \alpha)} = \lim_{t \rightarrow 0} \frac{\beta^3 e^{t\beta} - \alpha^3 e^{t\alpha}}{3(\beta - \alpha)} = \frac{\beta^3 - \alpha^3}{3(\beta - \alpha)} = \frac{\alpha^2 + \alpha\beta + \beta^2}{3} \\ V(X) &= \mu_2' - \mu^2 = \frac{\alpha^2 + \alpha\beta + \beta^2}{3} - \left(\frac{\alpha + \beta}{2}\right)^2 = \frac{(\beta - \alpha)^2}{12}\end{aligned}$$

5-81.

$$\begin{aligned}\text{a) } M_X(t) &= E(e^{tX}) = \int_0^\infty e^{tx} \lambda e^{-\lambda x} dx = \lambda \int_0^\infty e^{(t-\lambda)x} dx \\ &= \lambda \left(\frac{e^{(t-\lambda)x}}{t-\lambda} \right)_0^\infty \text{ which is finite only if } t < \lambda. \\ M_X(t) &= \lambda \left(\frac{e^{(t-\lambda)x}}{t-\lambda} \right)_0^\infty = \lambda \left(0 - \frac{1}{t-\lambda} \right) = \frac{\lambda}{\lambda-t} \quad \text{for } t < \lambda.\end{aligned}$$

$$\begin{aligned}\text{b) } E(X) &= \mu = \mu_1' = \frac{dM_X(t)}{dt} \Big|_{t=0} = \frac{\lambda}{(\lambda-t)^2} \Big|_{t=0} = \frac{1}{\lambda} \\ \mu_2' &= \frac{d^2M_X(t)}{dt^2} \Big|_{t=0} = \frac{2\lambda}{(\lambda-t)^3} \Big|_{t=0} = \frac{2}{\lambda^2} \\ V(X) &= \mu_2' - \mu^2 = \frac{2}{\lambda^2} - \left(\frac{1}{\lambda}\right)^2 = \frac{1}{\lambda^2}\end{aligned}$$

5-82.

$$\text{a) } M_X(t) = E(e^{tX}) = \int_0^\infty e^{tx} \frac{\lambda}{\Gamma(r)} (\lambda x)^{r-1} e^{-\lambda x} dx = \frac{\lambda^r}{\Gamma(r)} \int_0^\infty x^{r-1} e^{(t-\lambda)x} dx$$

$\int_0^\infty x^{r-1} e^{(t-\lambda)x} dx$ is finite only if $t < \lambda$. Besides, we need to use integration by substitution by

letting $z = (\lambda - t)x$. Note that the limits of the integration stay the same because $z \rightarrow 0$ as $x \rightarrow 0$, and $z \rightarrow \infty$ as $x \rightarrow \infty$. So, we have

$$\begin{aligned} x &= \frac{z}{(\lambda - t)} \quad \text{and} \quad dx = \frac{dz}{(\lambda - t)} \\ M_X(t) &= \frac{\lambda^r}{\Gamma(r)} \int_0^\infty x^{r-1} e^{(t-\lambda)x} dx = \frac{\lambda^r}{\Gamma(r)} \int_0^\infty \left(\frac{z}{\lambda - t}\right)^{r-1} e^{-z} \frac{dz}{\lambda - t} = \frac{\lambda^r}{\Gamma(r)(\lambda - t)^r} \int_0^\infty z^{r-1} e^{-z} dz \\ &= \frac{\lambda^r}{\Gamma(r)(\lambda - t)^r} \Gamma(r) = \frac{\lambda^r}{(\lambda - t)^r} = \left(\frac{\lambda - t}{\lambda}\right)^{-r} = \left(1 - \frac{t}{\lambda}\right)^{-r} \end{aligned}$$

As a result, $M_X(t) = \left(1 - \frac{t}{\lambda}\right)^{-r}$ for $t < \lambda$.

Also note that $\Gamma(r) = \int_0^\infty z^{r-1} e^{-z} dz$ for $r > 0$ by the definition of the gamma function.

$$\begin{aligned} \text{b) } E(X) &= \mu = \mu_1' = \frac{dM_X(t)}{dt} \Big|_{t=0} = \left(1 - \frac{t}{\lambda}\right)^{-r} \Big|_{t=0} = \lambda^r (\lambda - t)^{-r} \Big|_{t=0} = \lambda^r r (\lambda - t)^{-r-1} \Big|_{t=0} = \frac{r}{\lambda} \\ \mu_2' &= \frac{d^2M_X(t)}{dt^2} \Big|_{t=0} = r(r+1)\lambda^r (\lambda - t)^{-r-2} \Big|_{t=0} = \frac{r(r+1)}{\lambda^2} \\ V(X) &= \mu_2' - \mu^2 = \frac{r(r+1)}{\lambda^2} - \left(\frac{r}{\lambda}\right)^2 = \frac{r}{\lambda^2} \end{aligned}$$

5-83.

$$\text{a) } M_Y(t) = M_{X_1}(t)M_{X_2}(t)\dots M_{X_r}(t) = \frac{\lambda}{\lambda - t} \frac{\lambda}{\lambda - t} \dots \frac{\lambda}{\lambda - t} = \left(\frac{\lambda}{\lambda - t}\right)^r$$

$$\text{b) } M_Y(t) = \left(\frac{\lambda}{\lambda - t}\right)^r = \left(1 - \frac{t}{\lambda}\right)^{-r}$$

is the moment-generating function of a gamma distribution. As a result, the random variable Y has a gamma distribution with parameters r and λ .

5-84.

$$\text{a) } M_Y(t) = M_{X_1}(t)M_{X_2}(t) = \exp\left(\mu_1 t + \frac{\sigma_1^2 t^2}{2}\right) \times \exp\left(\mu_2 t + \frac{\sigma_2^2 t^2}{2}\right)$$

- $= \exp\left(\mu_1 t + \frac{\sigma_1^2 t^2}{2} + \mu_2 t + \frac{\sigma_2^2 t^2}{2}\right) = \exp\left((\mu_1 + \mu_2)t + (\sigma_1^2 + \sigma_2^2)\frac{t^2}{2}\right)$
- b) $M_Y(t) = \exp\left((\mu_1 + \mu_2)t + (\sigma_1^2 + \sigma_2^2)\frac{t^2}{2}\right)$ is the moment-generating function of a normal distribution with mean $\mu_1 + \mu_2$ and variance $\sigma_1^2 + \sigma_2^2$. As a result, the random variable Y has a normal distribution with parameters $\mu_1 + \mu_2$ and $\sigma_1^2 + \sigma_2^2$.

Supplemental Exercises

- 5-85. The sum of $\sum_x \sum_y f(x, y) = 1$, $\left(\frac{1}{4}\right) + \left(\frac{1}{8}\right) + \left(\frac{1}{8}\right) + \left(\frac{1}{4}\right) + \left(\frac{1}{4}\right) = 1$ and $f_{XY}(x, y) \geq 0$
- a) $P(X < 0.5, Y < 1) = f_{XY}(0,0) = 1/4$.
- b) $P(X \leq 1) = f_{XY}(0,0) + f_{XY}(0,1) + f_{XY}(1,0) + f_{XY}(1,1) = 3/4$
- c) $P(Y < 1.5) = f_{XY}(0,0) + f_{XY}(0,1) + f_{XY}(1,0) + f_{XY}(1,1) = 3/4$
- d) $P(X > 0.5, Y < 1.5) = f_{XY}(1,0) + f_{XY}(1,1) = 3/8$
- e) $E(X) = 0(3/8) + 1(3/8) + 2(1/4) = 7/8$
 $V(X) = 0^2(3/8) + 1^2(3/8) + 2^2(1/4) - 7/8^2 = 39/64$
 $E(Y) = 1(3/8) + 0(3/8) + 2(1/4) = 7/8$
 $V(Y) = 1^2(3/8) + 0^2(3/8) + 2^2(1/4) - 7/8^2 = 39/64$
- f) $f_X(x) = \sum_y f_{XY}(x, y)$ and $f_X(0) = 3/8$, $f_X(1) = 3/8$, $f_X(2) = 1/4$.
- g) $f_{Y|1}(y) = \frac{f_{XY}(1, y)}{f_X(1)}$ and $f_{Y|1}(0) = \frac{1/8}{3/8} = 1/3$, $f_{Y|1}(1) = \frac{1/4}{3/8} = 2/3$.
- h) $E(Y | X = 1) = \sum_{x=1} y f_{Y|X=1}(y) = 0(1/3) + 1(2/3) = 2/3$
- i) As is discussed in the chapter, because the range of (X, Y) is not rectangular, X and Y are not independent.
- j) $E(XY) = 1.25$, $E(X) = E(Y) = 0.875$, $V(X) = V(Y) = 0.6094$
 $\text{COV}(X, Y) = E(XY) - E(X)E(Y) = 1.25 - 0.875^2 = 0.4844$

$$\rho_{XY} = \frac{0.4844}{\sqrt{0.6094} \sqrt{0.6094}} = 0.7949$$

- 5-86. $P(X = 2, Y = 4, Z = 14) = \frac{20!}{2!4!14!} 0.05^2 0.25^4 0.70^{14} = 0.0385$
- b) $P(X = 0) = 0.05^0 0.95^{20} = 0.3585$
- c) $E(X) = np_1 = 20(0.05) = 1$
 $V(X) = np_1(1 - p_1) = 20(0.05)(0.95) = 0.95$
- d) $f_{X|Z=z}(X | Z = 19) \frac{f_{xz}(x, z)}{f_z(z)}$

$$f_{xz}(xz) = \frac{20!}{x!z!(20-x-z)!} 0.05^x 0.25^{20-x-z} 0.7^z$$

$$f_z(z) = \frac{20!}{z!(20-z)!} 0.3^{20-z} 0.7^z$$

$$f_{x|z=z}(X | Z=19) \frac{f_{xz}(x,z)}{f_z(z)} = \frac{(20-z)!}{x!(20-x-z)!} \frac{0.05^x 0.25^{20-x-z}}{0.3^{20-z}} = \frac{(20-z)!}{x!(20-x-z)!} \left(\frac{1}{6}\right)^x \left(\frac{5}{6}\right)^{20-x-z}$$

Therefore, X is a binomial random variable with $n=20-z$ and $p=1/6$. When $z=19$,

$$f_{x|19}(0) = \frac{5}{6} \text{ and } f_{x|19}(1) = \frac{1}{6}.$$

$$\text{e) } E(X | Z=19) = 0\left(\frac{5}{6}\right) + 1\left(\frac{1}{6}\right) = \frac{1}{6}$$

- 5-87. Let X, Y, and Z denote the number of bolts rated high, moderate, and low. Then, X, Y, and Z have a multinomial distribution.

$$\text{a) } P(X=12, Y=6, Z=2) = \frac{20!}{12!6!2!} 0.6^{12} 0.25^6 0.15^2 = 0.0422$$

b) Because X, Y, and Z are multinomial, the marginal distribution of Z is binomial with $n = 20$ and $p = 0.15$

$$\text{c) } E(Z) = np = 20(0.15) = 3$$

$$\text{d) } P(\text{low}>2) = 1 - P(\text{low}=0) - P(\text{low}=1) - P(\text{low}=2) =$$

$$\begin{aligned} & 1 - \frac{20!}{20!0!} (0.15)^0 (0.85)^{20} - \frac{20!}{19!1!} (0.15)^1 (0.85)^{19} - \frac{20!}{18!2!} (0.15)^2 (0.85)^{18} \\ & = 1 - 0.0388 - 0.1368 - 0.2293 = 0.5951 \end{aligned}$$

$$\text{e) } f_{z|16}(z) = \frac{f_{xz}(16,z)}{f_x(16)} \text{ and } f_{xz}(x,z) = \frac{20!}{x!z!(20-x-z)!} 0.6^x 0.25^{(20-x-z)} 0.15^z \text{ for } x+z \leq 20 \text{ and } 0 \leq x, 0 \leq z. \text{ Then,}$$

$$f_{z|16}(z) = \frac{\frac{20!}{16!z!(4-z)!} 0.6^{16} 0.25^{(4-z)} 0.15^z}{\frac{20!}{16!4!} 0.6^{16} 0.4^4} = \frac{4!}{z!(4-z)!} \left(\frac{0.25}{0.4}\right)^{4-z} \left(\frac{0.15}{0.4}\right)^z$$

for $0 \leq z \leq 4$. That is the distribution of Z given $X=16$ is binomial with $n=4$ and $p=0.375$

$$\text{f) From part (a), } E(Z) = 4(0.375) = 1.5$$

g) Because the conditional distribution of Z given $X=16$ does not equal the marginal distribution of Z, X and Z are not independent.

- 5-88. Let X, Y, and Z denote the number of calls answered in two rings or less, three or four rings, and five rings or more, respectively.

$$\text{a) } P(X=8, Y=1, Z=1) = \frac{10!}{8!1!1!} 0.7^8 0.20^1 0.10^1 = 0.1038$$

b) Let W denote the number of calls answered in four rings or less. Then, W is a binomial random variable with $n = 10$ and $p = 0.90$.

$$\text{Therefore, } P(W=10) = \binom{10}{10} 0.90^{10} 0.10^0 = 0.3487.$$

$$\text{c) } E(W) = 10(0.90) = 9.$$

$$\text{d) } f_{z|8}(z) = \frac{f_{xz}(8,z)}{f_x(8)} \text{ and } f_{xz}(x,z) = \frac{10!}{x!z!(10-x-z)!} 0.70^x 0.2^{(10-x-z)} 0.1^z \text{ for } x+z \leq 10 \text{ and } 0 \leq x, 0 \leq z. \text{ Then,}$$

$$f_{Z|8}(z) = \frac{\frac{10!}{8!z!(2-z)!} 0.70^8 0.2^{(2-z)} 0.1^z}{\frac{10!}{8!2!} 0.70^8 0.30^2} = \frac{2!}{z!(2-z)!} \left(\frac{0.2}{0.3}\right)^{2-z} \left(\frac{0.1}{0.3}\right)^z$$

for $0 \leq z \leq 2$. That is Z is binomial with $n=2$ and $p = 0.1/0.3 = 1/3$.

e) $E(Z)$ given $X = 8$ is $2(1/3) = 2/3$.

f) Because the conditional distribution of Z given $X = 8$ does not equal the marginal distribution of Z, X and Z are not independent.

5-89. $\int_0^3 \int_0^2 cx^2 y dy dx = \int_0^3 cx^2 \frac{y^2}{2} \Big|_0^2 dx = 2c \frac{x^3}{3} \Big|_0^3 = 18c$. Therefore, $c = 1/18$.

a) $P(X < 1, Y < 1) = \int_0^1 \int_0^1 \frac{1}{18} x^2 y dy dx = \int_0^1 \frac{1}{18} x^2 \frac{y^2}{2} \Big|_0^1 dx = \frac{1}{36} \frac{x^3}{3} \Big|_0^1 = \frac{1}{108}$

b) $P(X < 2.5) = \int_0^{2.5} \int_0^2 \frac{1}{18} x^2 y dy dx = \int_0^{2.5} \frac{1}{18} x^2 \frac{y^2}{2} \Big|_0^2 dx = \frac{1}{9} \frac{x^3}{3} \Big|_0^{2.5} = 0.5787$

c) $P(1 < Y < 2) = \int_0^3 \int_1^2 \frac{1}{18} x^2 y dy dx = \int_0^3 \frac{1}{18} x^2 \frac{y^2}{2} \Big|_1^2 dx = \frac{1}{12} \frac{x^3}{3} \Big|_0^3 = \frac{3}{4}$

d)

$$\begin{aligned} P(X > 2, 1 < Y < 1.5) &= \int_2^3 \int_1^{1.5} \frac{1}{18} x^2 y dy dx = \int_2^3 \frac{1}{18} x^2 \frac{y^2}{2} \Big|_1^{1.5} dx = \frac{5}{144} \frac{x^3}{3} \Big|_2^3 \\ &= \frac{95}{432} = 0.2199 \end{aligned}$$

e) $E(X) = \int_0^3 \int_0^2 \frac{1}{18} x^3 y dy dx = \int_0^3 \frac{1}{18} x^3 2 dx = \frac{1}{9} \frac{x^4}{4} \Big|_0^3 = \frac{9}{4}$

f) $E(Y) = \int_0^3 \int_0^2 \frac{1}{18} x^2 y^2 dy dx = \int_0^3 \frac{1}{18} x^2 \frac{8}{3} dx = \frac{4}{27} \frac{x^3}{3} \Big|_0^3 = \frac{4}{3}$

g) $f_X(x) = \int_0^2 \frac{1}{18} x^2 y dy = \frac{1}{9} x^2$ for $0 < x < 3$

h) $f_{Y|X}(y) = \frac{f_{XY}(1, y)}{f_X(1)} = \frac{\frac{1}{18} y}{\frac{1}{9}} = \frac{y}{2}$ for $0 < y < 2$.

i) $f_{X|1}(x) = \frac{f_{XY}(x, 1)}{f_Y(1)} = \frac{\frac{1}{18} x^2}{\frac{4}{27}} = \frac{27}{144} x^2$ and $f_Y(y) = \int_0^3 \frac{1}{18} x^2 y dx = \frac{y}{2}$ for $0 < y < 2$.

Therefore, $f_{X|1}(x) = \frac{\frac{1}{18} x^2}{1/2} = \frac{1}{9} x^2$ for $0 < x < 3$.

5-90. The region $x^2 + y^2 \leq 1$ and $0 < z < 4$ is a cylinder of radius 1 (and base area π) and height 4.

Therefore, the volume of the cylinder is 4π and $f_{XYZ}(x, y, z) = \frac{1}{4\pi}$ for $x^2 + y^2 \leq 1$ and $0 < z < 4$.

a) The region $X^2 + Y^2 \leq 0.5$ is a cylinder of radius $\sqrt{0.5}$ and height 4. Therefore,

$$P(X^2 + Y^2 \leq 0.5) = \frac{4(0.5\pi)}{4\pi} = 1/2.$$

b) The region $X^2 + Y^2 \leq 0.5$ and $0 < z < 1$ is a cylinder of radius $\sqrt{0.5}$ and height 1. Therefore,

$$P(X^2 + Y^2 \leq 0.5, Z < 1) = \frac{1(0.5\pi)}{4\pi} = 1/8$$

c) $f_{XY|1}(x, y) = \frac{f_{XYZ}(x, y, 1)}{f_Z(1)}$ and $f_Z(z) = \iint_{x^2+y^2 \leq 1} \frac{1}{4\pi} dy dx = 1/4$

for $0 < z < 4$. Then, $f_{XY|1}(x, y) = \frac{1/4\pi}{1/4} = \frac{1}{\pi}$ for $x^2 + y^2 \leq 1$.

d) $f_X(x) = \int_0^4 \int_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \frac{1}{4\pi} dy dz = \int_0^4 \frac{1}{2\pi} \sqrt{1-x^2} dz = \frac{2}{\pi} \sqrt{1-x^2}$ for $-1 < x < 1$

e) $f_{Z|0,0}(z) = \frac{f_{XYZ}(0,0,z)}{f_{XY}(0,0)}$ and $f_{XY}(x, y) = \int_0^4 \frac{1}{4\pi} dz = 1/\pi$ for $x^2 + y^2 \leq 1$. Then,

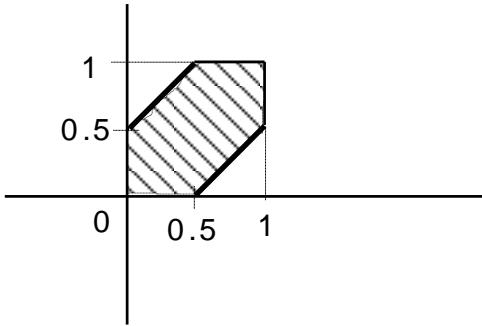
$$f_{Z|0,0}(z) = \frac{1/4\pi}{1/\pi} = 1/4 \text{ for } 0 < z < 4 \text{ and } \mu_{Z|0,0} = 2.$$

f) $f_{Z|xy}(z) = \frac{f_{XYZ}(x, y, z)}{f_{XY}(x, y)} = \frac{1/4\pi}{1/\pi} = 1/4$ for $0 < z < 4$. Then, $E(Z)$ given $X = x$ and $Y = y$ is

$$\int_0^4 \frac{z}{4} dz = 2.$$

5-91. $f_{XY}(x, y) = c$ for $0 < x < 1$ and $0 < y < 1$. Then, $\int_0^1 \int_0^1 c dx dy = 1$ and $c = 1$.

Because $f_{XY}(x, y)$ is constant, $P(|X - Y| < 0.5)$ is the area of the shaded region below



That is, $P(|X - Y| < 0.5) = 3/4$.

5-92. a) Let X_1, X_2, \dots, X_6 denote the lifetimes of the six components, respectively. Because of independence,

$$P(X_1 > 5000, X_2 > 5000, \dots, X_6 > 5000) = P(X_1 > 5000)P(X_2 > 5000)\dots P(X_6 > 5000)$$

If X is exponentially distributed with mean θ , then $\lambda = \frac{1}{\theta}$ and

$$P(X > x) = \int_x^{\infty} \frac{1}{\theta} e^{-t/\theta} dt = -e^{-t/\theta} \Big|_x^{\infty} = e^{-x/\theta}. \text{ Therefore, the answer is } e^{-1.25} e^{-1} e^{-0.5} e^{-0.4} = e^{-4.65} = 0.0096.$$

b) The probability that at least one component lifetime exceeds 25,000 hours is the same as 1 minus the probability that none of the component lifetimes exceed 25,000 hours. Thus,

$$1 - P(X_1 < 25,000, X_2 < 25,000, \dots, X_6 < 25,000) = 1 - P(X_1 < 25,000) \dots P(X_6 < 25,000) \\ = 1 - (1 - e^{-25/8})(1 - e^{-2.5})(1 - e^{-2.5})(1 - e^{-1.25})(1 - e^{-1.25})(1 - e^{-1}) = 1 - 0.2592 = 0.7408$$

- 5-93. Let X, Y, and Z denote the number of problems that result in functional, minor, and no defects, respectively.

a) $P(X = 2, Y = 5) = P(X = 2, Y = 5, Z = 3) = \frac{10!}{2!5!3!} 0.1^2 0.6^5 0.3^3 = 0.053$

b) Z is binomial with n = 10 and p = 0.3.

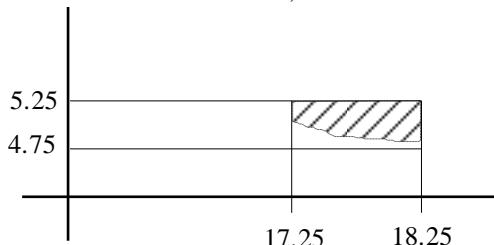
c) $E(Z) = 10(0.3) = 3$.

- 5-94. a) Let \bar{X} denote the mean weight of the 25 bricks in the sample. Then, $E(\bar{X}) = 1.5$ and

$$\sigma_{\bar{X}} = \frac{0.1}{\sqrt{25}} = 0.02. \text{ Then, } P(\bar{X} < 1.48) = P(Z < \frac{1.48-1.5}{0.02}) = P(Z < -1) = 0.159.$$

b) $P(\bar{X} > x) = P(Z > \frac{x-1.5}{0.02}) = 0.99. \text{ So, } \frac{x-1.5}{0.02} = -2.33 \text{ and } x = 2.9534.$

- 5-95. a) Because $\int_{17.75}^{18.25} \int_{4.75}^{5.25} c dy dx = 0.25c$, $c = 4$. The area of a panel is XY and $P(XY > 90)$ is the shaded area times 4 below,



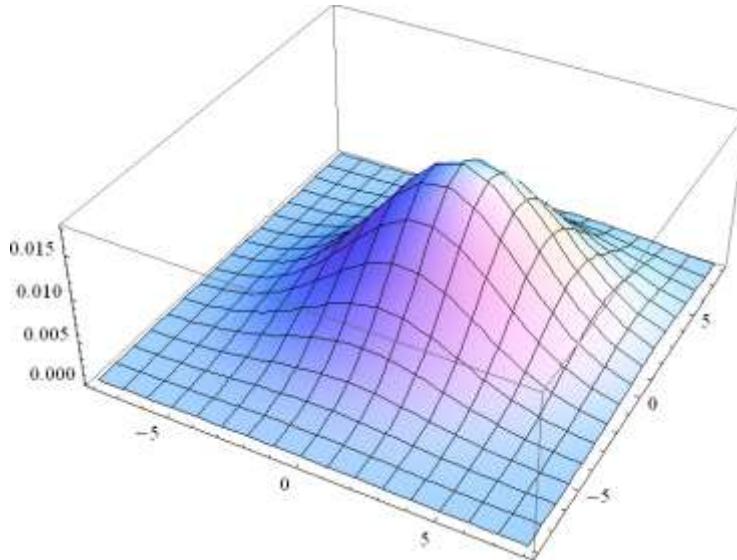
$$\text{That is, } \int_{17.75}^{18.25} \int_{90/x}^{5.25} 4 dy dx = 4 \int_{17.75}^{18.25} 5.25 - \frac{90}{x} dx = 4(5.25x - 90 \ln x) \Big|_{17.75}^{18.25} = 0.499$$

- b) The perimeter of a panel is $2X + 2Y$ and we want $P(2X + 2Y > 46)$

$$\int_{17.75}^{18.25} \int_{23-x}^{5.25} 4 dy dx = 4 \int_{17.75}^{18.25} 5.25 - (23 - x) dx \\ = 4 \int_{17.75}^{18.25} (-17.75 + x) dx = 4(-17.75x + \frac{x^2}{2}) \Big|_{17.75}^{18.25} = 0.5$$

- 5-96. a) Let X denote the weight of a piece of candy and $X \sim N(3, 0.3)$. Each package has 16 candies, then P is the total weight of the package with 16 pieces and $E(P) = 16(3) = 48$ g and $V(P) = 16^2 \times (0.3)^2 = 23.04$ g²
 b) $P(P < 48) = P(Z < \frac{48-48}{\sqrt{23.04}}) = P(Z < 0) = 0.5$.
 c) Let Y equal the total weight of the package with 17 pieces, $E(Y) = 17 \times (3) = 51$ g and $V(Y) = 17^2 \times (0.3)^2 = 26.01$ g²
 $P(Y < 1.6) = P(Z < \frac{48-51}{\sqrt{26.01}}) = P(Z < -0.59) = 0.2776$.
- 5-97. Let \bar{X} denote the average time to locate 15 parts. Then, $E(\bar{X}) = 45$ and $\sigma_{\bar{X}} = \frac{30}{\sqrt{15}}$
 a) $P(\bar{X} > 60) = P(Z > \frac{60-45}{30/\sqrt{15}}) = P(Z > 1.94) = 0.026$
 b) Let Y denote the total time to locate 15 parts. Then, $Y > 600$ if and only if $\bar{X} > 60$. Therefore, the answer is the same as part a.
- 5-98. a) Let Y denote the weight of an assembly. Then, $E(Y) = 4 + 5.5 + 10 + 8 = 27.5$ and $V(Y) = 0.4^2 + 0.5^2 + 0.2^2 + 0.5^2 = 0.7$.
 $P(Y > 29.5) = P(Z > \frac{29.5-27.5}{\sqrt{0.7}}) = P(Z > 2.39) = 0.0084$
 b) Let \bar{X} denote the mean weight of 8 independent assemblies. Then, $E(\bar{X}) = 27.5$ and $V(\bar{X}) = 0.7/8 = 0.0875$. Also, $P(\bar{X} > 29) = P(Z > \frac{29-27.5}{\sqrt{0.0875}}) = P(Z > 5.07) = 0$.

5-99.



5-100.

$$f_{XY}(x, y) = \frac{1}{1.2\pi} e^{\left[\frac{-1}{0.72} \{(x-1)^2 - 1.6(x-1)(y-2) + (y-2)^2\} \right]}$$

$$f_{XY}(x, y) = \frac{1}{2\pi\sqrt{.36}} e^{\left[\frac{-1}{2(0.36)} \{(x-1)^2 - 1.6(x-1)(y-2) + (y-2)^2\} \right]}$$

$$f_{XY}(x, y) = \frac{1}{2\pi\sqrt{1-.8^2}} e^{\left[\frac{-1}{2(1-0.8^2)} \{(x-1)^2 - 2(.8)(x-1)(y-2) + (y-2)^2\} \right]}$$

$$E(X) = 1, E(Y) = 2 \quad V(X) = 1 \quad V(Y) = 1 \quad \text{and} \quad \rho = 0.8$$

- 5-101. Let T denote the total thickness. Then, $T = X_1 + X_2$ and

a) $E(T) = 0.5 + 1 = 1.5 \text{ mm}$

$$V(T) = V(X_1) + V(X_2) + 2\text{Cov}(X_1 X_2) = 0.01 + 0.04 + 2(0.014) = 0.078 \text{ mm}^2$$

$$\text{where } \text{Cov}(XY) = \rho \sigma_X \sigma_Y = 0.7(0.1)(0.2) = 0.014$$

b) $P(T < 1.2) = P\left(Z < \frac{1.2 - 1.5}{\sqrt{0.078}}\right) = P(Z < -1.07) = 0.1423$

c) Let P denote the total thickness. Then, $P = 2X_1 + 3X_2$ and

$$E(P) = 2(0.5) + 3(1) = 4 \text{ mm}$$

$$V(P) = 4V(X_1) + 9V(X_2) + 2(2)(3)\text{Cov}(X_1 X_2) \\ = 4(0.01) + 9(0.04) + 2(2)(3)(0.014) = 0.568 \text{ mm}^2$$

$$\text{where } \text{Cov}(XY) = \rho \sigma_X \sigma_Y = 0.7(0.1)(0.2) = 0.014$$

- 5-102. Let T denote the total thickness. Then, $T = X_1 + X_2 + X_3$ and

a) $E(T) = 0.5 + 1 + 1.5 = 3 \text{ mm}$

$$V(T) = V(X_1) + V(X_2) + V(X_3) + 2\text{Cov}(X_1 X_2) + 2\text{Cov}(X_2 X_3) +$$

$$2\text{Cov}(X_1 X_3) = 0.01 + 0.04 + 0.09 + 2(0.014) + 2(0.03) + 2(0.009) = 0.246 \text{ mm}^2$$

$$\text{where } \text{Cov}(XY) = \rho \sigma_X \sigma_Y$$

b) $P(T < 1.4) = P\left(Z < \frac{1.4 - 3}{0.246}\right) = P(Z < -6.5) \approx 0$

- 5-103. Let X and Y denote the percentage returns for security one and two respectively.

If half of the total dollars is invested in each then $\frac{1}{2}X + \frac{1}{2}Y$ is the percentage return.

$$E(\frac{1}{2}X + \frac{1}{2}Y) = 0.05$$

$$V(\frac{1}{2}X + \frac{1}{2}Y) = 1/4 V(X) + 1/4 V(Y) + 2(1/2)(1/2)\text{Cov}(X, Y)$$

$$\text{where } \text{Cov}(XY) = \rho \sigma_X \sigma_Y = -0.5(2)(3) = -3$$

$$V(\frac{1}{2}X + \frac{1}{2}Y) = 1/4(4) + 1/4(9) - 1.5 = 1.75$$

$$\text{Also, } E(X) = 5 \text{ and } V(X) = 4.$$

Therefore, the strategy that splits between the securities has a lower standard deviation of percentage return than investing \$2 million in the first security.

- 5-104. a) The range consists of nonnegative integers with $x + y + z = 4$.

b) Because the samples are selected without replacement, the trials are not independent and the joint distribution is not multinomial.

c) $P(X = x | Y = 2) = \frac{f_{XY}(x, 2)}{f_Y(2)}$

$$P(Y = 2) = \frac{\binom{7}{0} \binom{5}{2} \binom{8}{2}}{\binom{20}{4}} + \frac{\binom{7}{1} \binom{5}{2} \binom{8}{1}}{\binom{20}{4}} + \frac{\binom{7}{2} \binom{5}{2} \binom{8}{0}}{\binom{20}{4}} = 0.05779 + 0.11558 + 0.04334 = 0.21671$$

$$P(X = 0 \text{ and } Y = 2) = \frac{\binom{7}{0} \binom{5}{2} \binom{8}{2}}{\binom{20}{4}} = 0.05779$$

$$P(X = 1 \text{ and } Y = 2) = \frac{\binom{7}{1} \binom{5}{2} \binom{8}{1}}{\binom{20}{4}} = 0.11558$$

$$P(X = 2 \text{ and } Y = 2) = \frac{\binom{7}{2} \binom{5}{2} \binom{8}{0}}{\binom{20}{4}} = 0.04334$$

x	$f_{XY}(x,2)$
0	$0.05779/0.21671 = 0.267$
1	$0.11558/0.21671 = 0.535$
2	$0.04334/0.21671 = 0.19999$

d)

$P(X=x, Y=y, Z=z)$ is the number of subsets of size 4 that contain x printers with graphics enhancements, y printers with extra memory, and z printers with both features divided by the number of subsets of size 4.

$$P(X = x, Y = y, Z = z) = \frac{\binom{7}{x} \binom{5}{y} \binom{8}{z}}{\binom{20}{4}} \quad \text{for } x + y + z = 4.$$

$$P(X = 1, Y = 2, Z = 1) = \frac{\binom{7}{1} \binom{5}{2} \binom{8}{1}}{\binom{20}{4}} = 0.11558$$

$$\text{e)} \ P(X = 1, Y = 1) = P(X = 1, Y = 1, Z = 2) = \frac{\binom{7}{1} \binom{5}{1} \binom{8}{2}}{\binom{20}{4}} = 0.2023$$

f) The marginal distribution of X is hypergeometric with $N = 20$, $n = 4$, $K = 4$. Therefore, $E(X) = nK/N = 4/5$ and $V(X) = 4(4/20)(16/20)(16/19) = 0.5389$.

$$\text{g)} \ P(X = 1, Y = 2 | Z = 1) = P(X = 1, Y = 2, Z = 1) / P(Z = 1)$$

$$= \left[\frac{\binom{7}{1} \binom{5}{2} \binom{8}{1}}{\binom{20}{4}} \right] / \left[\frac{\binom{8}{1} \binom{12}{3}}{\binom{20}{4}} \right] = 0.3181$$

$$\text{h)} \ P(X = 2 | Y = 2) = P(X = 2, Y = 2) / P(Y = 2)$$

$$= \left[\frac{\binom{7}{2} \binom{5}{1} \binom{8}{0}}{\binom{20}{4}} \right] / \left[\frac{\binom{5}{2} \binom{15}{2}}{\binom{20}{4}} \right] = 0.2$$

i) Because $X + Y + Z = 4$, if $Y = 0$ and $Z = 3$, then $X = 1$. Because X must equal 1, $f_{X|YZ}(1) = 1$.

- 5-105. a) Let X , Y , and Z denote the risk of new competitors as no risk, moderate risk, and very high risk. Then, the joint distribution of X , Y , and Z is multinomial with $n = 12$ and $p_1 = 0.15$, $p_2 = 0.70$, and $p_3 = 0.15$. X , Y and $Z \geq 0$ and $x + y + z = 12$
- b) $P(X = 1, Y = 3, Z = 1) = 0$, not possible since $x + y + z \neq 12$

$$\text{c)} \ P(Z \leq 2) = \binom{12}{0} 0.15^0 0.85^{12} + \binom{12}{1} 0.15^1 0.85^{11} + \binom{12}{2} 0.15^2 0.85^{10}$$

$$= 0.1422 + 0.3012 + 0.2924 = 0.7358$$

$$\text{d)} \ P(Z = 2 | Y = 1, X = 10) = 0$$

$$\text{e)} \ P(X = 10) = P(X = 10, Y = 2, Z = 0) + P(X = 10, Y = 1, Z = 1) + P(X = 10, Y = 0, Z = 2)$$

$$= \frac{12!}{10!2!0!} 0.15^{10} 0.70^2 0.15^0 + \frac{12!}{10!1!1!} 0.15^{10} 0.70^1 0.15^1 + \frac{12!}{10!0!2!} 0.15^{10} 0.70^0 0.15^2$$

$$= 1.86 \times 10^{-7} + 7.99 \times 10^{-8} + 8.56 \times 10^{-9} = 2.745 \times 10^{-7}$$

$$\begin{aligned}
 P(Z \leq 1 | X = 10) &= \frac{P(Z = 0, Y = 2, X = 10)}{P(X = 10)} + \frac{P(Z = 1, Y = 1, X = 10)}{P(X = 10)} \\
 &= \frac{12!}{10!2!0!} 0.15^{10} 0.70^2 0.15^0 / 2.745 \times 10^{-7} + \frac{12!}{10!1!1!} 0.15^{10} 0.70^1 0.15^1 / 2.745 \times 10^{-7} \\
 &= 0.9687
 \end{aligned}$$

$$\begin{aligned}
 \text{f)} \quad P(Y \leq 1, Z \leq 1 | X = 10) &= \frac{P(Z = 1, Y = 1, X = 10)}{P(X = 10)} \\
 &= \frac{12!}{10!1!1!} 0.15^{10} 0.70^1 0.15^1 / 6.89 \times 10^{-8} \\
 &= 0.2912
 \end{aligned}$$

$$\begin{aligned}
 \text{g)} \quad E(Z | X = 10) &= (0(1.86 \times 10^{-7}) + 1(7.99 \times 10^{-8}) + 2(8.56 \times 10^{-9}) / 2.745 \times 10^{-7}) \\
 &= 0.353
 \end{aligned}$$

Mind-Expanding Exercises

5-106. By the independence,

$$\begin{aligned}
 P(X_1 \in A_1, X_2 \in A_2, \dots, X_p \in A_p) &= \int_{A_1} \int_{A_2} \dots \int_{A_p} f_{X_1}(x_1) f_{X_2}(x_2) \dots f_{X_p}(x_p) dx_1 dx_2 \dots dx_p \\
 &= \left[\int_{A_1} f_{X_1}(x_1) dx_1 \right] \left[\int_{A_2} f_{X_2}(x_2) dx_2 \right] \dots \left[\int_{A_p} f_{X_p}(x_p) dx_p \right] \\
 &= P(X_1 \in A_1) P(X_2 \in A_2) \dots P(X_p \in A_p)
 \end{aligned}$$

5-107. $E(Y) = c_1\mu_1 + c_2\mu_2 + \dots + c_p\mu_p$. Also,

$$\begin{aligned}
 V(Y) &= \int [c_1x_1 + c_2x_2 + \dots + c_px_p - (c_1\mu_1 + c_2\mu_2 + \dots + c_p\mu_p)]^2 f_{X_1}(x_1) f_{X_2}(x_2) \dots f_{X_p}(x_p) dx_1 dx_2 \dots dx_p \\
 &= \int [c_1(x_1 - \mu_1) + \dots + c_p(x_p - \mu_p)]^2 f_{X_1}(x_1) f_{X_2}(x_2) \dots f_{X_p}(x_p) dx_1 dx_2 \dots dx_p
 \end{aligned}$$

Now, the cross-term

$$\begin{aligned}
 &\int c_1 c_2 (x_1 - \mu_1)(x_2 - \mu_2) f_{X_1}(x_1) f_{X_2}(x_2) \dots f_{X_p}(x_p) dx_1 dx_2 \dots dx_p \\
 &= c_1 c_2 \left[\int (x_1 - \mu_1) f_{X_1}(x_1) dx_1 \right] \left[\int (x_2 - \mu_2) f_{X_2}(x_2) dx_2 \right] = 0
 \end{aligned}$$

from the definition of the mean. Therefore, each cross-term in the last integral for $V(Y)$ is zero and

$$\begin{aligned}
 V(Y) &= \left[\int c_1^2 (x_1 - \mu_1)^2 f_{X_1}(x_1) dx_1 \right] \dots \left[\int c_p^2 (x_p - \mu_p)^2 f_{X_p}(x_p) dx_p \right] \\
 &= c_1^2 V(X_1) + \dots + c_p^2 V(X_p).
 \end{aligned}$$

$$\begin{aligned}
 \text{5-108. } \int_0^a \int_0^b f_{XY}(x, y) dy dx &= \int_0^a \int_0^b c dy dx = cab. \text{ Therefore, } c = 1/ab. \text{ Then, } f_X(x) = \int_0^b c dy = \frac{1}{a} \\
 &\text{for } 0 < x < a, \text{ and } f_Y(y) = \int_0^a c dx = \frac{1}{b} \text{ for } 0 < y < b. \text{ Therefore, } f_{XY}(x, y) = f_X(x)f_Y(y) \text{ for all } x \\
 &\text{and } y \text{ and } X \text{ and } Y \text{ are independent.}
 \end{aligned}$$

5-109. The marginal density of X is

$$f_X(x) = \int_0^b g(x)h(u)du = g(x) \int_0^b h(u)du = kg(x) \text{ where } k = \int_0^b h(u)du. \text{ Also,}$$

$$f_Y(y) = lh(y) \text{ where } l = \int_0^a g(v)dv. \text{ Because } f_{XY}(x,y) \text{ is a probability density function,}$$

$$\int_0^a \int_0^b g(x)h(y)dydx = \left[\int_0^a g(v)dv \right] \left[\int_0^b h(u)du \right] = 1. \text{ Therefore, } kl = 1 \text{ and}$$

$$f_{XY}(x,y) = f_X(x)f_Y(y) \text{ for all } x \text{ and } y.$$

$$5-110. \text{ The probability function for } X \text{ is } P(X = x) = \frac{\binom{m}{k} \binom{N-m}{n-k}}{\binom{N}{n}}$$

The number of ways to select x_j items from N_j is $\binom{N_j}{x_j}$.

Therefore, from the multiplication rule the total number of ways to select items to meet the

$$\text{conditions is } \binom{N_1}{x_1} \binom{N_2}{x_2} \dots \binom{N_k}{x_k}$$

The total number of subsets of n items selected from N is $\binom{N}{n}$. Therefore

$$P(X_1 = x_1, \dots, X_k = x_k) = \frac{\binom{N_1}{x_1} \binom{N_2}{x_2} \dots \binom{N_k}{x_k}}{\binom{N}{n}}$$