

§ 4.7

Normal Approximation to the Binomial and Poisson Distributions

If X is binomial random variable

$$\text{Mean} = np$$

$$\sigma^2 = np(1-p)$$

→ approximate to normal $\Rightarrow Z = \frac{X - \mu}{\sigma} = \frac{X - np}{\sqrt{np(1-p)}}$

* Continuity correction: (± 0.5)

$$P(X \leq x) = P(X \leq x + 0.5)$$

$$P(X \geq x) = P(X \geq x - 0.5)$$

- ① Example: Assume that in a digital communication channel, the number of bits received in error can be modeled by a binomial random variable, and assume that the probability that a bit is received in error is 10^{-5} . If 16 million bits are transmitted, what is the probability that 150 or fewer errors occur?

$$P = 10^{-5} \quad n = 16 \text{ million} = 16 \times 10^6 \quad P(X \leq 150) = ?$$

1- By binomial $\sum_{x=0}^{150} \binom{16 \times 10^6}{x} (10^{-5})^x (1 - 10^{-5})^{16 \times 10^6 - x}$
approximation \rightarrow $\text{cumulative } X \text{ is small}$
 $x=0 \rightarrow x=150 \text{ (approx)} \leq 150$

2- By Normal Approximation:

$$P(X \leq 150) = P(X \leq 150 + 0.5) = P(X \leq 150.5)$$

$$P(Z \leq \frac{150.5 - (16 \times 10^6)(10^{-5})}{\sqrt{(16 \times 10^6)(10^{-5})(1 - 10^{-5})}}) = P(Z \leq \frac{150.5 - 160}{\sqrt{12.649}})$$

$$P(Z \leq -0.751) = 0.227 \rightarrow \text{From table}$$

* normal approximation is work at n large relative to p .
as the previous example.

* binomial approximation is effective if
 $np > 5, n(1-p) > 5$

$\frac{n}{N} < 0.1 \rightarrow n = \text{sample size}$
 $N = \text{population size}$

Normal approximation for Poisson random variable:

$$Z = \frac{X - \lambda}{\sqrt{\lambda}} \Rightarrow \lambda = E(x) = V(x)$$

* Using the same correction used for binomial (± 0.5)

** The approximation is good for $\lambda > 5$

② Example: Assume that the number of asbestos particles in a squared meter of dust on a surface follows a Poisson distribution with a mean of 1000. If a squared meter of dust is analyzed, what is the probability that 950 or fewer particles are found?

$$P(X \leq 950)$$

$$P(X \leq 950) = P(X \leq 950 + 0.5) = P(X \leq 950.5)$$

$$= P\left(Z \leq \frac{950.5 - 1000}{\sqrt{1000}}\right) = P(Z \leq -1.57) = 0.058208$$

§ 24.8 Exponential Distributions

Exponential random variable: X = is equals the distance between successive events from a Poisson process with mean number of events $\lambda > 0$ per unit interval, with parameter λ .

The probability density function of X is:

$$f(x) = \lambda e^{-\lambda x} \quad 0 \leq x < \infty$$

• این توزیع از توزیع Poisson process وابسته است. Exponential dist. ***

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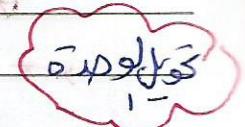
$$E(x) = \mu = \frac{1}{\lambda} \text{ (mean)}$$

$$V(x) = \sigma^2 = \frac{1}{\lambda^2}$$

process (so) distribution
→ Exponential

③ Example: In a large corporate computer network, user log-ons to the system can be modeled as a Poisson process with a mean of 25 log-ons per hour. What is the probability that there are no log-ons in an interval of six minutes?

no log-ons \Rightarrow failure (failure)

(X) 6 minutes \Rightarrow $X > 6 \text{ minutes} \Rightarrow X > 0.1 \text{ hour} \Rightarrow$ 

$$P(0.1 < X) = \int_{-\infty}^{\infty} 1 e^{-\lambda x} dx = \int_{-\infty}^{\infty} 25 e^{-25x} dx$$
$$= \left[-e^{-25x} \right]_{0.1}^{\infty} = -e^{-25x} \Big|_{0.1}^{\infty} = e^{-25x} \Big|_{0.1}^{\infty} = e^{-2.5} - e^{-\infty} = e^{-2.5} - 0$$

as $x \rightarrow \infty, e^{-25x} \rightarrow 0$

$$P(0.1 < X) = 0.0821$$

② * What is the probability that the next time until the next log-on is between 2 & 3 minutes?

$$P(2 \text{ min} < X < 3 \text{ min}) = P(0.033 < X < 0.05)$$

$$\cancel{=} \int_{0.033}^{0.05} 1 e^{-\lambda x} dx = \int_{0.033}^{0.05} 25 e^{-25x} dx$$
$$= e^{-25x} \Big|_{0.033}^{0.05} = e^{-25 \times 0.033} - e^{-25 \times 0.05} = 0.438 - 0.287$$
$$= 0.1517$$

③ * Determine the interval of time such that the probability that no log-on occurs in the interval is 0.9.

time = x in minutes.

$$P(X > x) = e^{-25x} = 0.9$$

$$-25x = \ln 0.9$$

$$x = 4.214 \times 10^{-3} \text{ (hours)} \quad | \quad 60 \text{ min}$$

1 hour

$$x = 0.25 \text{ minute}$$

④ * The mean time until the next log-on & its standard deviation:

$$\mu = \sigma^2 = \frac{1}{\lambda} = \frac{1}{25} = 0.04 \text{ hour} = 2.4 \text{ minutes}$$

Ch #5: Joint Probability Distributions:

§ 5.1: Two or More Random variables:
joint probability distribution:

The probability distribution that defines X & Y (random variables) & their simultaneous behavior

bivariate probability distribution: The joint probability distribution of 2 random variables, is usually written as $P(X=x, Y=y)$
discrete

joint probability mass function: $f_{XY}(x,y)$ satisfies:

$$(1) f_{XY}(x,y) \geq 0$$

$$(2) \sum_x \sum_y f_{XY}(x,y) = 1$$

$$(3) f_{XY}(x,y) = P(X=x, Y=y)$$

* probability outside the range $(X, Y) = 0$
" " range $x = 0$

Joint probability density function: (volume under the surface $f_{XY}(x,y)$)
statifies the following properties:

$$(1) f_{XY}(x,y) \geq 0 \text{ for all } x, y$$

$$(2) \iint_{-\infty}^{\infty} f_{XY}(x,y) dx dy = 1$$

$$(3) P((X,Y) \in R) = \iint_R f_{XY}(x,y) dx dy$$

R = region of 2D space

Examples: X: the time until a computer server connects to the machine
 Y: the time until the server authorizes you as a valid user

Assume that the joint probability function for X & Y :

$$f_{XY}(x,y) = 6x^{10^{-6}} \exp(-0.001x - 0.002y) \quad x < y$$

① Show that $f_{XY}(x,y)$ is density function.

$\int_{-\infty}^{\infty} f_{XY}(x,y) dy = 1$ if $f_{XY}(x,y)$ is density function. $f_{XY}(x,y)$ \leftarrow what?

$y = x$ \leftarrow أفق $\leftarrow x < y$ if x \leftarrow في الأقوان حكماً *
 $y = \infty$ \leftarrow أفق

بما أنه عدد صران أحصى بالخطام للرسوت عد كيمن أذ يكون

$$\text{أفق} \leftarrow x=0 \quad \leftarrow \text{ع}$$

و بما أن y بدلالة x ببراء التكامل بـ (dy)

$$\int_0^\infty \int_x^\infty 6 \times 10^{-6} \exp(-0.001x - 0.002y) dy dx$$

$$6 \times 10^{-6} \int_0^\infty \left[\exp(-0.001x) \cdot \boxed{\exp(-0.002y)} dy \right] dx$$

$$\frac{6 \times 10^{-6}}{-0.002} \int_0^\infty \exp(-0.001x) [\exp(-0.002 \cdot 0) - \exp(-0.002x)] dx$$

$$3 \times 10^{-3} \int_0^\infty \exp(-0.001x) \cdot \exp(-0.002x) dx$$

$$3 \times 10^{-3} \int_0^\infty \exp(-0.003x) dx = \frac{3 \times 10^{-3}}{-0.003} [\exp(-0.003 \cdot 0) - \exp(-0.003x)]$$

$$= -1 \times -1 = \boxed{1} \quad \text{X} \neq 1$$

② The probability that $X < 1000$ & $Y < 2000$?

$$x=0 \quad \text{origin} \quad x=1000 \quad \text{point} \leftarrow X$$

$$y=x \quad \text{line} \quad y=2000 \quad \text{line} \leftarrow Y$$

$$6 \times 10^{-6} \int_0^{1000} \int_x^{2000} \exp(-0.001x) \exp(-0.002y) dy dx$$

$$\frac{6 \times 10^{-6}}{-0.002} \int_{1000}^\infty \exp(-0.001x) [\exp(-0.002 \cdot 2000) - \exp(-0.002x)] dx$$

$$3 \times 10^{-3} \int_0^\infty \exp(y) \exp(-0.001x) - \exp(-0.003x) dx$$

$$3 \times 10^{-3} \left[\frac{\exp(-4)}{-0.001} [\exp(-0.001 \cdot 1000) - \exp(0)] - \left[\frac{\exp(-0.003 \cdot 1000) - \exp(0)}{-0.003} \right] \right]$$

$$3 \times 10^{-3} [305.16] = 0.915$$

Marginal Probability Distributions:

The individual probability distribution of a random variable that can be determined from the joint probability distribution

marginal probability mass function:

$$f_X(x) = \sum_{\substack{y=1 \\ \text{small}}}^{\infty} f_{xy}(x,y)$$

⑤ Examples

$y = \text{Response Time}$	$x = \text{Number of Bars of Single Strength}$			Marginal of y
	1	2	3	
4	0.15	0.1	0.05	0.3
3	0.02	0.1	0.05	0.17
2	0.02	0.03	0.2	0.25
1	0.01 0.2	0.02 0.25	0.25 0.55	0.28
	Marginal of x			

$$f_X(3) = P(X=3) = \sum_{x=3} \sum_{y=1}^{x,y} f_{xy}$$

$$= P(X=3, Y=1) + P(X=3, Y=2) + P(X=3, Y=3) + P(X=3, Y=4)$$

$$= 0.25 + 0.2 + 0.05 + 0.05$$

Marginal probability density function:

$$f_X(x) = \int_{-\infty}^b f_{xy}(x,y) dy$$

$$f_Y(y) = \int_a^{\infty} f_{xy}(x,y) dx$$

$$P(a < x < b) = \int_a^b \int_{-\infty}^b f_{xy}(x,y) dy dx$$

Conditional Probability Distribution:

$$P(Y=B | X=A) = \frac{P(X=A, Y=B)}{P(X=A)} \quad \left. \begin{array}{l} \\ \end{array} \right\} \text{Mass function}$$

Example: (S continue)

$$(1) P(Y=1 | X=3) = \frac{P(X=3, Y=1)}{P(X=3)} = \frac{0.25}{0.55} = 0.454$$

↑ Marginal

(2) The probability that $Y=2$ given that $X=3$:

$$P(Y=2 | X=3) = \frac{P(X=3, Y=2)}{P(X=3)} = \frac{0.2}{0.55} = \frac{f_{XY}(3, 2)}{f_X(3)} = 0.364$$

$$(3) P(Y=1 | X=3) + P(Y=2 | X=3) + P(Y=3 | X=3) + P(Y=4 | X=3)$$

↑ Marginal

$$= \frac{0.25}{0.55} + \frac{0.2}{0.55} + \frac{0.05}{0.55} + \frac{0.05}{0.55} = \frac{0.55}{0.55} = [1]$$

Conditional Probability density function:

$$f_{Y|X}(y) = \frac{f_{XY}(x, y)}{f_X(x)} \quad \text{for } f_X(x) \geq 0$$

Properties:

$$(1) f_{Y|X}(y) \geq 0$$

$$(2) \int f_{Y|X}(y) dy = 1$$

$$(3) P(Y \in B | X=x) = \int_B f_{Y|X}(y) dy \quad B \rightarrow \text{range of } Y$$

Example (1) continue:

$$f_{xy}(x,y) = 6 \times 10^{-6} \exp(-0.001x - 0.002y)$$

(1) Marginal density function of x for $x > 0$

(y) \rightarrow infinity \leftarrow Marginal for (x) w/p a_1 in

$$y = x \quad \text{and} \quad y > x \quad ; \quad y \rightarrow \infty$$

$$f_x(x) = \int_x^{\infty} 6 \times 10^{-6} \exp(-0.001x) \exp(-0.002y) dy$$

$$6 \times 10^{-6} \exp(-0.001x) \int_x^{\infty} \exp(-0.002y) dy$$

$$\frac{6 \times 10^{-6} \exp(-0.001x)}{-0.002} \left[\exp(-0.002 \cdot \cancel{x}) - \exp(-0.002x) \right]^0$$

$$f_x(x) = 3 \times 10^{-3} \exp(-0.003x)$$

(2) conditional probability density function:

$$f_{Y|X}(y) = \frac{f_{XY}(x,y)}{f_X(x)} = \frac{6 \times 10^{-6} \exp(-0.001x) \exp(-0.002y)}{3 \times 10^{-3} \exp(-0.003x)}$$

$$= 2 \times 10^{-3} \exp(-0.002y) \exp(0.002x)$$

(3) The conditional probability density function of y , given that $X = 1500$, is non zero. Determine the

probability that Y exceeds 2000, given that $x=1500$

$$P(Y > 2000 | x=1500) = \int_{2000}^{\infty} f_{Y|x=1500}(y) dy$$

$$= \int_{2000}^{\infty} 2 \times 10^{-3} \exp(0.002 \overset{(1500)}{*}) \exp(-0.002 y) dy$$

$$2 \times 10^{-3} \exp(0.002 \overset{(1500)}{*}) \int_{2000}^{\infty} \exp(-0.002 y) dy$$

$$\frac{2 \times 10^{-3}}{-0.002} \exp(0.002 \overset{(1500)}{*}) \left[\exp(-0.002 \overset{0}{*}) - \exp(-0.002 \overset{-4}{*} 2000) \right]$$
$$- \exp(3) * -\exp(-4) = 0.368$$

Conditional mean & variance:

Conditional mean: $E(Y|x) = \mu_{Y|x} = \cancel{\int y P_{Y|x}(y) dy} \int y f_{Y|x}(y)$

conditional variance: $V(Y|x) = \sigma_{Y|x}^2 = \int y^2 f_{Y|x}(y) dy - (\mu_{Y|x})^2$

Independence

If X & Y are independent $\Rightarrow f_{XY}(x,y) = f_X(x)f_Y(y)$

$$(1) f_{XY}(x,y) = f_X(x)f_Y(y) \rightarrow \text{for all } x \in \mathbb{R}, y \in \mathbb{R}$$

$$(2) f_{Y|X}(y|x) = f_Y(y) \rightarrow \text{for all } x \in \mathbb{R}, y \in \mathbb{R} \text{ with } f_X(x) > 0$$

$$(3) f_{X|Y}(x|y) = f_X(x) \rightarrow \text{for all } x \in \mathbb{R}, y \in \mathbb{R} \text{ with } f_Y(y) > 0$$

$$(4) P(X \in A, Y \in B) = P(X \in A)P(Y \in B) \rightarrow \text{for any } A, B \text{ in the range } X, Y$$

example

$$f_{XY}(x,y) = 2 \times 10^{-6} \exp(-0.002y) \exp(-0.001x)$$

$$x \geq 0 \quad \& \quad y \geq 0$$

(1) Show that X & Y are independent.

$$f_{XY}(x,y) = f_X(x)f_Y(y)$$

$$f_X(x) = \int_0^\infty [2 \times 10^{-6} \exp(-0.002y) \exp(-0.001x)] dy$$

$$= \frac{2 \times 10^{-6}}{0.002} \exp(-0.001x) [\exp(0) - \exp(-\infty)] = 10^{-3} \exp(-0.001x)$$

$$f_Y(y) = \int_0^\infty 2 \times 10^{-6} \exp(-0.002y) \exp(-0.001x) dx$$

$$= 2 \times 10^{-3} \exp(-0.002y) [\exp(-0.001 \times 0) - \exp(-0.001 \times \infty)]$$

$$= 2 \times 10^{-3} \exp(-0.002y)$$

$$\Rightarrow 2 \times 10^{-3} \exp(-0.002y) * 10^{-3} \exp(-0.001x)$$

$$= 2 \times 10^{-6} \exp(-0.002y) \exp(-0.001x) = f_{XY}(x,y)$$

(2) Determine $P(X > 1000, Y < 1000)$

$$P(X > 1000, Y < 1000) = P(X > 1000) P(Y < 1000)$$

$$= \int_{1000}^{\infty} f_X(x) dx * \int_0^{1000} f_Y(y) dy$$

$$= \int_{1000}^{\infty} 10^{-3} \exp(-0.001x) dx * \int_0^{1000} 2 \times 10^{-3} \exp(-0.002y) dy$$

$$= \frac{10}{0.001} [\exp(-1) - \exp(-\infty)] * \frac{2 \times 10^{-3}}{0.002} (\exp(0) - \exp(-2))$$

$$\exp(-1) * (1 - \exp(-2)) = 0.318$$

Example: Let the random variables X & Y denote the lengths of two dimensions of a machined part, respectively. Assume that X & Y are independent random variables, and further assume that the distribution of X is normal with mean 10.5 mm & variance 0.0025 mm 2 and that distribution of Y is normal with mean 3.2 mm & variance 0.0036 . Determine the probability that $10.4 < X < 10.6$ & $3.15 < Y < 3.25$.

$$P(10.4 < X < 10.6, 3.15 < Y < 3.25) = ??$$

$$= P(10.4 < X < 10.6) * P(3.15 < Y < 3.25)$$

$$= P\left(\frac{10.4 - 10.5}{\sqrt{0.0025}} < Z < \frac{10.6 - 10.5}{\sqrt{0.0025}}\right) * P\left(\frac{3.15 - 3.2}{\sqrt{0.0036}} < Z < \frac{3.25 - 3.2}{\sqrt{0.0036}}\right)$$

$$P(-2 < Z < 2) * P(-0.83 < Z < 0.83)$$

$$= \{P(Z < 2) - P(Z < -2)\} * \{P(Z < 0.83) - P(Z < -0.83)\}$$

$$= 0.9542 * 0.5937 = 0.567$$

§ 5.2: Covariance & Correlation & Association

Expected Values

$$E[h(x, Y)] = \begin{cases} \sum h(x, y) f_{xy}(x, y) & \rightarrow X, Y \text{ Discrete} \\ \int \int h(x, y) f_{xy}(x, y) dx dy & \rightarrow X, Y \text{ continuous} \end{cases}$$

Example: (5 continue)

For the joint probability distribution of the 2 random variables calculate $E[(X - M_x)(Y - M_y)]$

$$M_x = 2.35$$

$$M_y = 2.49$$

X	Y	$X - M_x$	$Y - M_y$	$(X - M_x)(Y - M_y)$	$(2 - M_x)(Y - M_y)$	$(3 - M_x)(Y - M_y)$
1	1	-1.35	-1.49	2.0115	0.5215	0.9685
2	2	-0.35	-0.49	0.6615	0.1715	-0.3185
3	3	0.65	0.51	-0.6885	-0.1785	0.3315
4			1.51	-2.0385	-0.5285	0.9815

$$\begin{aligned} \sum \sum (X - M_x)(Y - M_y) f_{xy}(x, y) &= 2.0115 * 0.01 + 0.6615 * 0.02 \\ &+ -0.6885 * 0.02 + -2.0385 * 0.15 + 0.5215 * 0.02 + 0.6615 * 0.03 \\ &+ -0.1785 * 0.1 - 0.5285 * 0.1 - 0.9685 * 0.25 - 0.3185 * 0.2 + 0.3315 * 0.05 \\ &+ 0.05 * 0.9815 \\ &= -0.5815 \end{aligned}$$

Covariance: $\text{Cov}(X, Y)$ OR σ_{xy}

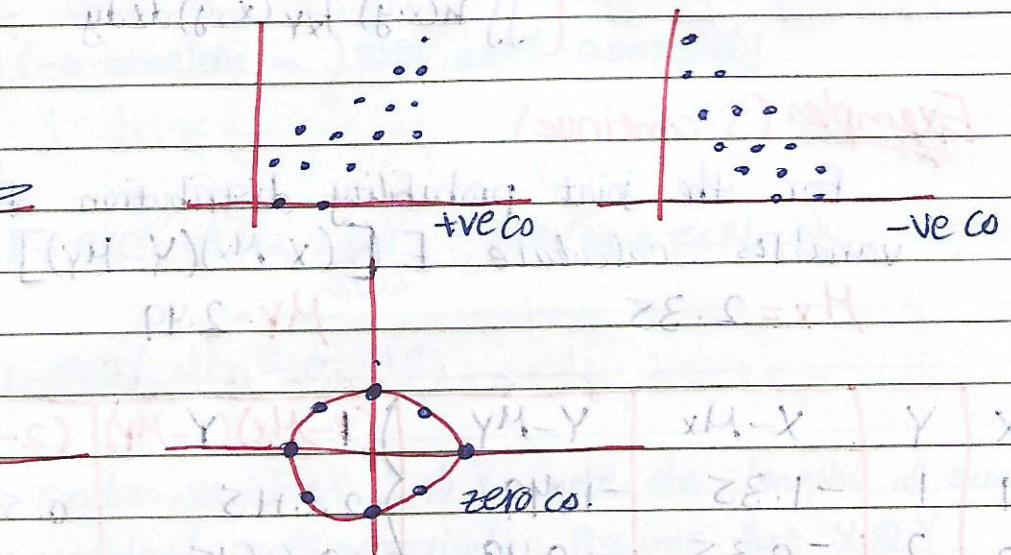
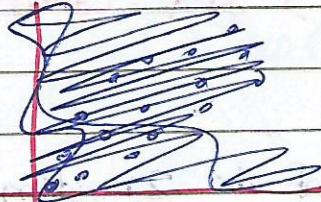
$$\sigma_{xy} = E((X - M_x)(Y - M_y)) = E(XY) - M_x M_y$$

Covariance is a measure of linear relationship between the random variables

* if the relationship between random variables is non linear?

the covariance might not be sensitive to the relationship

** Assumes all points are equally likely $E(Y|X)$



$$\text{Correlation } \rho_{xy} = \frac{\text{cov}(x, y)}{\sqrt{\text{var}(x)\text{var}(y)}} = \frac{\sigma_{xy}}{\sigma_x \sigma_y}$$

$$-1 \leq \rho_{xy} \leq 1$$

Correlation: is a dimensionless quantity that can be used to compare the linear relationships between pairs of variables in different units.

* 2 random variables with non zero correlation are said to be correlated.

** the correlation is a measure of the linear relationship between random variables.

$$\text{Cov}(Y, X) = E[(Y - \bar{Y})(X - \bar{X})] = E[(Y_i - \bar{Y})(X_i - \bar{X})] = \bar{Y}\bar{X}$$

Examples

Determine Cov_{XY} , P_{XY}

$$\text{Cov}_{XY} = E(XY) - E(X)E(Y)$$

$$E(XY) = \sum \sum XY f_{XY}(x,y) \rightarrow \text{discrete}$$

$$= 0*0*0.2 + 1*1*0.1 + 1*2*0.1$$

$$+ 1*1*0.1 + 2*2*0.1 + 0.4*3*0.4 = 0.1 + 0.2 + 0.2 + 0.4 + 3.6$$

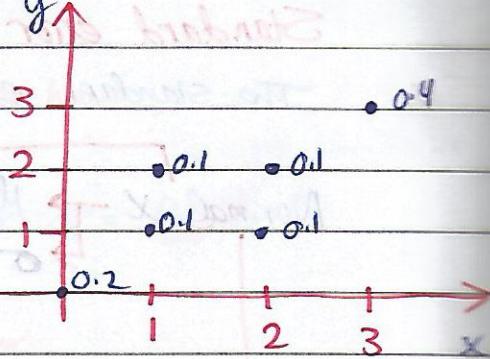
$$E(XY) = 4.5$$

$$E(X) = \sum X f_{XY}(x,y) = 0*0.2 + 0.2*1 + 0.7*2 + 3*0.4$$

$$= 0 + 0.2 + 0.4 + 1.2 = 1.8$$

$$E(Y) = \sum Y f_{XY}(x,y) = 0*0.2 + 1*0.2 + 2*0.2 + 3*0.4 = 1.8$$

$$\text{Cov}_{XY} = 4.5 - 1.8 * 1.8 = 4.5 - 3.24 = 1.26$$



$$(2) P_{XY} = \frac{\text{Cov}_{XY}}{\sigma_X \sigma_Y}$$

$$\sigma_X^2 = \sum (X - E(X))^2 f_{XY}(x,y) = (0-1.8)^2 * 0.2 + (1-1.8)^2 * 0.2 + (2-1.8)^2 * 0.2 + (3-1.8)^2 * 0.4 = 0.648 + 0.128 + 0.008 + 0.576$$

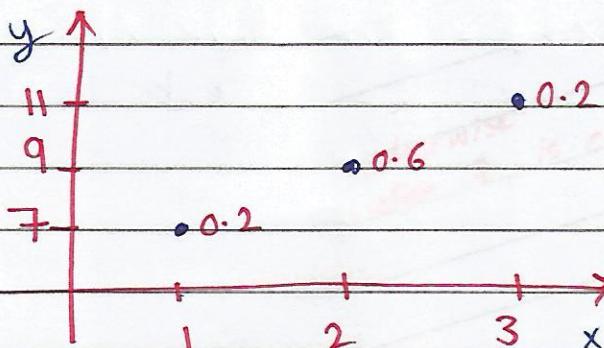
$$\sigma_X = 1.36$$

$$\sigma_Y^2 = \sum (Y - \mu_Y)^2 f_{XY}(x,y) = (0-1.8)^2 * 0.2 + (1-1.8)^2 * 0.2 + (2-1.8)^2 * 0.2 + (3-1.8)^2 * 0.4 = 1.36$$

$$P_{XY} = \frac{1.26}{\sqrt{1.36} \sqrt{1.36}} = 0.926$$

Examples Suppose that the random variable X has the following distributions $P(X=1)=0.2$, $P(X=2)=0.6$, $P(X=3)=0.2$. Let $Y = 2X + 5$. That is $P(Y=7)=0.2$ $P(Y=9)=0.6$.

Determine the correlation between X & Y .



$$Y = 2X + 5 \text{ about } Y = 7$$

$$x=1 \Rightarrow Y=7$$

$$x=2 \Rightarrow Y=9$$

$$x=3 \Rightarrow Y=11$$

Ch 7 & Point Estimation of Parameters

Sampling distributions

§ 7.1: Point Estimation

* **Sample distribution**: the probability distribution of a statistic

* ~~Statistic~~ is a random variable, such as: mean, variance, ...

* **the objective of point estimation**: to select a single number based on sample data that is the most plausible value for θ .

* **point estimate**: numerical value $\hat{\theta}$ of a sample ~~statistic~~ ~~with best unbiased~~ Statistic

$\Rightarrow \hat{\theta}$ = point estimate (numerical value) ~~is derived from sample~~

$\hat{\theta}$ = point estimator (sample statistic $\rightarrow h(X_1, X_2, \dots)$) ~~which is used~~

§ 7.2: Sampling Distributions & the Central Limit Theorem

* The random variables X_1, X_2, \dots, X_n are a random sample of size n if:

(a) the X_i 's are independent random variables

(b) every X_i 's has the same probability distribution

* the primary purpose in taking a random sample is to obtain information about the unknown population parameters.

* **Statistic**: any function of the observations in a random sample.

* Sampling of a statistic depends on a population in

1- distribution

2- the size of the sample

3- the method of sample selection.

Sampling distribution of the mean.

$$\bar{X}(\text{point estimator}) = \frac{x_1 + x_2 + \dots + x_n}{n}$$

$$M\bar{x} \text{ (normal distribution)} = \frac{\mu + \mu + \mu + \dots + \mu}{n} = \mu$$

$$\sigma_{\bar{x}}^2 = \frac{\sigma^2 + \sigma^2 + \dots + \sigma^2}{n} = \frac{\sigma^2}{n} \Rightarrow \sigma_{\bar{x}} = \frac{\sigma}{\sqrt{n}}$$

Example 8 An ~~electronics~~ company manufactures resistors that have a mean resistance of $\mu = 100$ ohms & a standard deviation of 10 ohms. The distribution of resistance is normal. Find the probability that a random sample of $n = 25$ resistors will have an average resistance of fewer than 95 ohms. $P(\bar{x} < 95)$

$$\begin{aligned} P(\bar{x} < 95) &= P\left(z < \frac{95 - 100}{\frac{10}{\sqrt{25}}}\right) = P\left(z < \frac{-5}{2}\right) = P(z < -2.5) \\ &= 0.0062 \text{ (rare event)} \end{aligned}$$

Central limit theorem

If X_1, X_2, \dots, X_n is a random sample of size n taken from a population (finite/infinite) with mean μ and finite variance σ^2 & if \bar{X} is a sample mean, the limiting form of the distribution of

$$Z = \frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}}$$

is the standard normal distribution

* $n > 30 \rightarrow$ the central limit theorem will always apply.

$$M_{\bar{X}_1 - \bar{X}_2} = M_{\bar{X}_1} - M_{\bar{X}_2} = \mu_1 - \mu_2$$

$$\sigma_{\bar{X}_1 - \bar{X}_2}^2 = \sigma_{\bar{X}_1}^2 + \sigma_{\bar{X}_2}^2 = \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}$$

IF we have 2 independent populations with means μ_1 & μ_2 and variances σ_1^2 & σ_2^2 and if \bar{X}_1 & \bar{X}_2 are the sample means of 2 independent random samples of size n_1 & n_2 from these populations, then the sampling distribution of

$$Z = \frac{\bar{X}_1 - \bar{X}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$$

is approximately standard normal if the conditions of the central limit theorem apply.

* If 2 populations are normal, the sampling distribution of Z is exactly standard normal.

Example: random variables:

\bar{X}_1 is normal with $\mu_1 = 5000 \quad \sigma_1 = 40 \quad n_1 = 16$

\bar{X}_2 is normal with $\mu_2 = 5050 \quad \sigma_2 = 30 \quad n_2 = 25$

What is the probability that the difference in the 2 samples means $\bar{X}_2 - \bar{X}_1$ is at least 25?

$$P(\bar{X}_2 - \bar{X}_1 \geq 25) = P\left(\frac{\bar{X}_2 - \bar{X}_1 - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}} \geq \frac{25 - (5000 - 5050)}{\sqrt{\frac{(40)^2}{16} + \frac{(30)^2}{25}}}\right)$$

$$P\left(Z \geq \frac{-25}{\sqrt{136}}\right) = P(Z \geq -2.14) = 0.9838$$

§ 7-3.3: Standard error & Reporting a Point Estimate

The Standard error of an estimator $\hat{\theta}$ is its standard deviation given by $\sigma_{\hat{\theta}} = \sqrt{J(\hat{\theta})}$. If the standard error involves unknown parameters that can be estimated, substitution of those values into $\sigma_{\hat{\theta}}$ produces an estimated standard error, denoted by $\hat{se}(\hat{\theta})$.

$$\sigma_{\bar{X}} = \frac{\sigma}{\sqrt{n}} \rightarrow \text{Standard error of } \bar{X}$$

$$SE(\bar{X}) = \hat{\sigma}_{\bar{X}} = \frac{s}{\sqrt{n}} \rightarrow \text{estimated standard error of } \bar{X}$$

\nwarrow Population σ is unknown
 \uparrow Population s is known
 Population is sample $\rightarrow s$ is known

Examples Data - (sample)

41.60 41.48 42.34 41.95 41.86
42.18 41.72 42.26 41.81 42.04

$$(1) \bar{x} = \frac{\sum x}{n} = \frac{419.24}{10} = 41.924 \rightarrow \text{sample mean}$$

$$(2) \sigma_{\bar{x}} = \frac{\sigma}{\sqrt{n}} \quad \sigma \text{ unknown for population}$$

But $s = 0.284$

$$SE(\bar{x}) = \hat{\sigma}_{\bar{x}} = \frac{s}{\sqrt{n}} = \frac{0.284}{\sqrt{10}} = 0.0898$$

Ch 8: Statistical Intervals for a Single Sample

§ 8.1.1: Development of the confidence interval § Its Basic properties.

Confidence interval: interval of the form ~~$\mu \in [L, U]$~~

$$P\{L \leq \mu \leq U\} = 1 - \alpha$$

$L \rightarrow$ lower-confidence limit $U \rightarrow$ upper-confidence limit

$1 - \alpha \rightarrow$ confidence coefficient where $0 < \alpha < 1$

standardize ~~μ~~

$$\rightarrow P\left(-z_{\frac{\alpha}{2}} < \frac{\bar{x} - \mu}{\sigma/\sqrt{n}} < z_{\frac{\alpha}{2}}\right) = 1 - \alpha$$

$$\rightarrow P\left\{\bar{x} - z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}} < \mu < \bar{x} + z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}\right\} = 1 - \alpha$$

random interval

If \bar{x} is the ~~mean~~ sample mean of a random sample of size n from a normal population with known variance σ^2 ,

a $100(1 - \alpha)\%$ CI (confidence interval) on μ is given by

$$\bar{x} - z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}} \leq \mu \leq \bar{x} + z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}$$

where $z_{\frac{\alpha}{2}}$ is the upper $100\alpha\%$ of the standard σ^2 normal distribution

Examples-

64.1 64.7 64.5 64.6 64.5 64.3 64.6 64.8
64.2 64.3

$$\sigma = 1$$

Find 95% CI

$$\bar{x} = \frac{\sum x}{n} = \frac{644.6}{10} = 64.46$$

$$n=10 \quad \sigma=1$$

$$95\% = 100(1-\alpha) \times$$

$$1-\alpha = 0.95$$

$$\Rightarrow \alpha = 0.05$$

$$z_{\frac{\alpha}{2}} = z_{0.025} = z_{0.025} = 1.96 \Rightarrow$$

بروح على $z_{0.025}$ هو 1.96 وبما أن $\alpha = 0.025$ يعني $z_{0.025}$ الذي يحيط بالوسط

$$\bar{x} - z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}} \leq M \leq \bar{x} + z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}$$

$$64.46 - 1.96 * \frac{1}{\sqrt{10}} \leq M \leq 64.46 + 1.96 * \frac{1}{\sqrt{10}}$$

$$63.84 \leq M \leq 65.08$$